Elevated Platelet Count Appears to Be Causally Associated with Increased Risk of Lung Cancer: A Mendelian Randomization Analysis

Ying Zhu1, Yongyue Wei1,2,3, Ruyang Zhang1,2,3, Xuesi Dong4, Silpeng Shen2, Yang Zhao1,3, Jianling Bai1, Demetrius Albanés5, Neil E. Caporaso6, Maria Teresa Landi7, Bin Zhu5, Stephen J. Chanock5, Fangyi Gu6, Stephen Lam7, Ming-Sound Tsao8, Frances A. Shepherd8, Adonina Tardon9, Ana Fernández-Somoano9, Guillermo Fernández-Tardon9, Chu Chen10, Matthew J. Barnett10, Jennifer Doherty10, Stig E. Bojesen11,12,13, Mattias Johansson14, Paul Brennan14, James D. McKay14, Robert Carreras-Torres14, Thomas Muley15,16, Angela Risch16,17,18, Heunz-Erich Wichmann19, Heike Bickeboeller20, Albert Rosenberger20, Gad Rennert21, Walid Saliba21, Susanne M. Arnold22, John K. Field23, Michael P.A. Davies23, Michael W. Marcus23, XiFeng Wu24, Yuanying Ye24, Loic Le Marchand25, Lynne R. Wilkens25, Olle Melander26, Jonas Manjer26, Hans Brunstrom27, Rajayen J. Hung28, Geoffrey Liu28, Yonathan Brhane28, Linda Kachuri28, Angelina S. Andrew29, Eric J. Duel30, Lambertus A. Kiemeney30, Erik HFM van der Heijden31, Aage Haugen32, Shanbeh Zienoddiny32, Vidar Kvaag32, Kjell Granvist33, Mikael Johansson33, Penella J. Woll34, Angela Cox34, Fiona Taylor34, Dawn M. Teare35, Philip Lazarus36, Matthew B. Schabath37, Melinda C. Aldrich38, Richard S. Houlston39, John McLaughlin40, Victoria L. Stevens41, Hongbing Shen42, Xifeng Wu24, Yuanqing Ye24, Fangyi Gu6, Stephen Lam7, Ming-Sound Tsao8, Steven J. Chanock5, Fangyi Gu6, Stephen Lam7, Ming-Sound Tsao8, Gary E. Goodman44, Feng Chen1,3, and David C. Christiani1,2,3

Note: Supplementary data for this article are available at Cancer Epidemiology, Biomarkers & Prevention Online (http://cebp.aacrjournals.org/).

Y. Zhu, Y. Wei, and R. Zhang contributed equally to this article.

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Research Article

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Abstract

Background: Platelets are a critical element in coagulation and inflammation, and activated platelets are linked to cancer risk through diverse mechanisms. However, a causal relationship between platelets and risk of lung cancer remains unclear.

Methods: We performed single and combined multiple instrumental variable Mendelian randomization analysis by an inverse-weighted method, in addition to a series of sensitivity analyses. Summary data for associations between SNPs and platelet count are from a recent publication that included 48,666 Caucasian Europeans, and the International Lung Cancer Consortium and Transdisciplinary Research in Cancer of the Lung study. The lung was recently noted to play a major role in platelet biogenesis and as an ideal bioreactor for production of mature platelets from megakaryocytes, which account for approximately 50% of total platelet production. Platelets are an important element in coagulation and inflammation, and diverse mechanisms link platelets to cancer progression. It has been identified that several variants in those chromosomal regions associated with platelet count have associations with myocardial infarction and autoimmune and hematologic disorders. Tumor-educated blood platelets have emerged as promising biomarkers for noninvasive detection of cancer, and it was demonstrated to discriminate patients with non–small cell lung cancer (NSCLC) from healthy individuals and patients with various nonsmoking inflammatory conditions. High platelet count is associated with increased mortality in a variety of cancers, including malignant mesothelioma, gynecologic malignancies, and breast cancer. In addition, platelet-to-lymphocyte ratio and mean platelet volume also add value in early diagnosis of lung cancer and prognosis prediction. These findings, taken together, indicate that disordered platelet production may be connected to lung carcinogenesis. However, due to potential unmeasured confounders in observational studies, the association between platelet count and lung cancer risk remains unclear.

Mendelian randomization is based on the principle that an individual's genotype is randomized at conception and utilizes genetic variants as instrumental variables for the association between phenotypic exposures and outcomes to eliminate bias due to unmeasured confounders. Genetic variables used as instrumental variables should meet the following assumptions: (i) genetic variants are associated with exposure, (ii) genetic variants affect outcome only via the exposure, and (iii) genetic variants are not associated with any confounders of the exposure–outcome association. Mendelian randomization analysis has been broadly used to estimate unconfounded associations between exposure and outcome, such as the effect of higher adult height on escalated cancer risk.

Results: Multiple instrumental variable analysis incorporating six SNPs showed a 62% increased risk of overall non–small cell lung cancer [NSCLC; OR, 1.62; 95% confidence interval (CI), 1.15–2.27; P = 0.005] and a 200% increased risk for small-cell lung cancer (OR, 3.00; 95% CI, 1.27–7.06; P = 0.01). Results showed only a trending association with NSCLC histologic subtypes, which may be due to insuffcient sample size and/or weak effect size. A series of sensitivity analysis retained these findings.

Conclusions: Our findings suggest a causal relationship between elevated platelet count and increased risk of lung cancer and provide evidence of possible antiplatelet interventions for lung cancer prevention.

Impact: These findings provide a better understanding of lung cancer etiology and potential evidence for antiplatelet interventions for lung cancer prevention.

Introduction

Lung cancer, a highly invasive, rapidly metastasizing cancer, has been the leading cause of cancer-related deaths worldwide for decades, accounting for more than one million deaths each year. Smoking is a major risk factor for lung cancer and for decades, accounting for more than one million deaths each year (1). In addition, environmental–occupational exposures (2), lifestyle, and genetic variants (3) have been broadly explored as risks/predisposing factors for lung cancer. However, aspects of lung cancer risk remain largely unexplained and thus warrant further study.

The lung was recently noted to play a major role in platelet biogenesis and act as an ideal bioreactor for production of mature platelets from megakaryocytes, which account for approximately 50% of total platelet production (6). Platelets are an important element in coagulation and inflammation, and diverse mechanisms link platelets to cancer progression (7, 8). It has been identified that several variants in those chromosomal regions associated with platelet count have associations with myocardial infarction and autoimmune and hematologic disorders. Tumor-educated blood platelets have emerged as promising biomarkers for noninvasive detection of cancer, and it was demonstrated to discriminate patients with non–small cell lung cancer (NSCLC) from healthy individuals and patients with various nonsmoking inflammatory conditions (9, 10). Indeed, high platelet count is associated with increased mortality in a variety of cancers, including malignant mesothelioma (11), gynecologic malignancies (12), and breast cancer (13). In addition, platelet-to-lymphocyte ratio and mean platelet volume also add value in early diagnosis of lung cancer (14) and prognosis prediction (15, 16). These findings, taken together, indicate that disordered platelet production may be connected to lung carcinogenesis. However, due to potential unmeasured confounders in observational studies, the association between platelet count and lung cancer risk remains unclear.

Mendelian randomization is based on the principle that an individual's genotype is randomized at conception (17) and utilizes genetic variants as instrumental variables for the association between phenotypic exposures and outcomes to eliminate bias due to unmeasured confounders. Genetic variants used as instrumental variables should meet the following assumptions: (i) genetic variants are associated with exposure, (ii) genetic variants affect outcome only via the exposure, and (iii) genetic variants are not associated with any confounders of the exposure–outcome association (18). By finding a genetic marker that satisfies instrumental variable assumptions, Mendelian randomization analysis has been broadly used to estimate unconfounded associations between exposure and outcome, such as the effect of higher adult height on escalated cancer risk (20–24).

In this study, we performed summary data–based Mendelian randomization (25) analysis, which is the extension of two sample Mendelian randomization, using curated platelet count–related SNPs as instrumental variables to evaluate the association between platelet count and lung cancer risk by using summary statistics from recent large-scale genome-wide association studies (GWAS).

Materials and Methods

Data source and study population

Mendelian randomization analysis was conducted to estimate the effect of platelet count (X) on risk of lung cancer (Y) using genetic variants (G) as instrumental variables (26). According to the Mendelian randomization analysis diagram described in Fig. 1, we used coefficients of genetic variants on platelet count (βpx), and their standard errors (SEpx) from the recently published study of Gieger and colleagues, which pooled 23 studies and included approximately 48,666 individuals of European descent (27).

A total of 54 genetic variants were identified that were associated with platelet count (Supplementary Table S1). One of the key assumptions underlying Mendelian randomization is that the genetic variants (SNPs) used as instrumental variables are only related to the outcome of interest through the exposure variable under study. No pleiotropic pathways should exist from platelet-related SNPs to lung cancers through intermediates other than platelet count. Thus, six genetic variants (rs17030845, rs6141, rs3792366, rs210134, rs708382, and rs6065) where further related SNPs as instrumental variables to evaluate the association between platelet count and lung cancer risk by using summary statistics from recent large-scale genome-wide association studies (GWAS).
platelets. By the way, the SNP rs6141 in THPO narrowly misses the level required for nominal significance ($P < 5 \times 10^{-8}$) with $P = 6.18 \times 10^{-8}$ in Europeans, but shows genome-wide significance in Japanese (28). Therefore, it is still included serving as instrumental variable for platelet count.

Coefficients ($b_{YX}$) and corresponding standard errors ($SE_{YX}$) of the association between genetic variants and lung cancer risk were obtained from meta-analysis of existing OncoArray and TRICL GWAS studies, which were detailed previously (29). Briefly, overall NSCLC samples were composed from OncoArray and TRICL GWASs, including 29,266 cases and 56,450 controls, and subgroup analyses were performed for 11,273 adenocarcinoma, 7,426 squamous cell carcinoma (SCC), and 2,664 small-cell lung cancer (SCLC) cases (Supplementary Table S2).

Mendelian randomization analysis

Mendelian randomization analysis with multiple instrumental variables was performed using an inverse-variance weighted (IVW) method combining the effect of genetic variants by weighted score. This score was used as an instrumental variable to estimate the effect of platelet count on lung cancer risk (26):

$$ \hat{b}_{YX_{IVW}} = \frac{\sum N_i b_{YX_i} SE_{YX_i}}{\sum N_i SE_{YX_i}^2} $$

$$ SE_{YX_{IVW}} = \sqrt{\frac{1}{\sum N_i SE_{YX_i}^2}} $$

In which $N = 6$ represents the number of instrumental variables included, and $\hat{b}_{YX_{IVW}}$ and $SE_{YX_{IVW}}$ represent the effect of platelet count on lung cancer risk in log(OR) scale and its corresponding SE. Associations of platelet count on risk of overall NSCLC and individual subtypes were analyzed. Results are presented as OR for lung cancer risk per $100 \times 10^9$/L increment of platelet count.

In addition, penalized IVW, robust IVW, MR-Egger, penalized MR-Egger, and robust MR-Egger methods were used for sensitivity analyses to evaluate robustness of the findings (30). Step forward modeling was used to add an optimal instrumental variable each time from the left 48 SNPs, adding to the six curated SNPs for multiple instrumental variable analysis, until there was no improvement of statistical significance ($P$) for the test of causal effect. The modeling process was terminated when no added SNP increased $-\log_{10}(P)$ by 20% or 10%. Besides, Mendelian randomization analysis with a single-instrumental variable (one SNP at a time) was performed as supplementary. Effect of platelet count on lung cancer risk ($b_{YX}$ in logOR scale) and its standard error ($SE_{YX}$) were estimated as follows (31):

$$ \hat{b}_{YX} = \frac{b_{YX_i} SE_{YX_i}}{SE_{YX_{IVW}}} $$

All analyses were performed using R Software Version 3.3.1 (The R Foundation). All tests were two-sided, and $P \leq 0.05$ was considered statistically significant unless stated otherwise.

Results

Among 48,666 Europeans, 54 SNPs were quantitatively associated with platelet count with $P \leq 5 \times 10^{-8}$ (Supplementary Table S1; ref. 27). Associations of those 54 SNPs with risk of lung cancer were analyzed among 29,266 cases and 56,450 controls from OncoArray and previous GWAS studies. Demographics and study descriptions were detailed previously (29) and are briefly listed in Supplementary Table S2 as well. Summarized association results of SNPs and lung cancer risk are listed in Supplementary Table 1.

Table 1. SNPs of specific platelet-related genes

<table>
<thead>
<tr>
<th>SNP</th>
<th>Chrposition (hg19)</th>
<th>Gene</th>
<th>Reference allele</th>
<th>Effect allele</th>
<th>EAF (%)</th>
<th>Function</th>
<th>b (95% CI)</th>
<th>P</th>
</tr>
</thead>
<tbody>
<tr>
<td>rs7030845</td>
<td>2: 43687879</td>
<td>THADA</td>
<td>C</td>
<td>T</td>
<td>9.65</td>
<td>Intron</td>
<td>-3.58 (-4.67 to -2.49)</td>
<td>1.27 × 10^{-10}</td>
</tr>
<tr>
<td>rs6141</td>
<td>3: 184090266</td>
<td>THPO</td>
<td>T</td>
<td>C</td>
<td>47.39</td>
<td>3' UTR</td>
<td>-2.47 (-3.36 to -1.57)</td>
<td>6.18 × 10^{-8}</td>
</tr>
<tr>
<td>rs3792566</td>
<td>6: 122839876</td>
<td>PDIAS</td>
<td>A</td>
<td>G</td>
<td>38.68</td>
<td>Intron</td>
<td>2.533 (1.44-2.87)</td>
<td>3.60 × 10^{-9}</td>
</tr>
<tr>
<td>rs210134</td>
<td>6: 33540209</td>
<td>BAKI</td>
<td>G</td>
<td>A</td>
<td>29.29</td>
<td>500 bp Downstream</td>
<td>-4.96 (-5.73 to -4.18)</td>
<td>7.11 × 10^{-5}</td>
</tr>
<tr>
<td>rs708382</td>
<td>7: 42442544</td>
<td>FAM171A2-</td>
<td>T</td>
<td>C</td>
<td>39.66</td>
<td>2 kb Upstream</td>
<td>-2.44 (-3.28 to -1.59)</td>
<td>1.51 × 10^{-4}</td>
</tr>
<tr>
<td>rs66055</td>
<td>17: 4856381</td>
<td>GPIBA</td>
<td>C</td>
<td>T</td>
<td>8.53</td>
<td>Missense</td>
<td>4.19 (2.96-5.43)</td>
<td>2.92 × 10^{-11}</td>
</tr>
</tbody>
</table>

Abbreviations: CI, confidence interval; EAF, effect allele frequency; UTR, untranslated region.
Table S3. According to instrumental variable assumptions that had evidence only related to platelets, 6 SNPs which are relatively independent and situated in different chromosomes were selected for Mendelian randomization analysis (Table 1), and 48 SNPs were excluded (Supplementary Table S4).

In multiple instrumental variable analysis combining all six relatively independent SNPs situated in different chromosomes, a significant association between platelet count and overall NSCLC risk is revealed, showing that each $100 \times 10^9/L$ increment of platelet count was associated with a 62% increase in NSCLC risk [95% confidence interval (CI), 1.15–2.27; $P = 0.005$; Figs. 2A and 3A]. In addition, five different methods of sensitivity analysis, including penalized IVW, robust IVW, MR-Egger, penalized MR-Egger, and robust MR-Egger, retained this association (Table 2). In NSCLC subtype analysis, it failed to detect significant associations between platelet count and the risk of lung adenocarcinoma (OR, 1.51; 95% CI, 0.92–2.48; $P = 0.11$; Figs. 2B and 3B) and SCC (OR, 1.59; 95% CI, 0.86–2.92; $P = 0.14$; Figs. 2C and 3C). On the other hand, it is suggested that platelet count is significantly associated with the risk of SCLC (OR, 3.00; 95% CI, 1.27–7.06; $P = 0.01$; Figs. 2D and 3D). The results of single-instrumental variable are presented in Supplementary Table S5. No correction was conducted for them because a single-weak instrument will have lower power to reject the null hypothesis (32).

We also performed a step forward modeling strategy to include more instrumental SNPs in the multiple instrumental variable model. Including more SNPs as instrumental variables yielded similar, yet more significant, causal estimates (Supplementary Table S6; Supplementary Fig. S1).

Discussion

This Mendelian randomization study suggests that each $100 \times 10^9/L$ increment in platelets results in a 62% increased risk of NSCLC and, notably, a 200% increased risk of SCLC. However, this study failed to show evidence of a relationship between

Figure 2.

Causal associations between platelet count and lung cancer risk. Forest plots of causal associations between platelet count (PLT) and risk of lung cancer using Mendelian randomization analysis incorporating different genetic variants as instrumental variables (IVs). Associations of platelet count with risk of NSCLC (A), adenocarcinoma (AC; B), SCC (C), and SCLC (D) were analyzed based on single-instrumental variable or multiple instrumental variables using IVW analysis.

Figure 3.

Associations between SNPs and lung cancer risk. Scatter plots displaying estimates of the association between each SNP and risk of lung cancer against quantitative relationship of each SNP on platelet count for NSCLC (A), adenocarcinoma (AC; B), SCC (C), and SCLC (D). Slope of the gray dashed line through the plot represents IVW regression estimate for the causal effect of platelet count on lung cancer risk.
platelet count and risk of adenocarcinoma and SCC, probably resulting from insufficient sample size. As comparing with SCLC, the effect size of platelet count on adenocarcinoma and SCC are weaker, larger sample size is needed (33).

Platelets have been studied for decades as an important regulator of inflammation and thrombosis (34), which are broadly interrelated with human carcinogenesis (13). Platelets are also recognized as a stimulator of proangiogenic factors (13) and a major source of VEGF (35), platelet-derived growth factor (36), and basic fibroblast growth factor (37), which act as promoters of tumor growth in lung (38–44). New evidence suggests that platelets are relevant to defensive, physiologic immune responses of the lungs and to inflammatory lung diseases (45). Thus, higher platelet count has a potential biological connection to increased risk of lung cancer. Interestingly, p-selectin, an important adhesion molecule expressed on the surface of activated platelets, is more highly expressed in lung adenocarcinomas and SCC than in healthy populations (46). These results indicate a considerable role of platelets in lung carcinogenesis.

Intriguingly, a recent study indicates that cancer cells depend on platelets to avoid anoikis and succeed in metastasis (47). Platelets induce resistance to anoikis in vitro and are critical for metastasis in vivo by activating RhoA-MYPT1-PP1–mediated YAP1 dephosphorylation and promoting its nuclear translocation to inhibit apoptosis. However, the unknown underlying mechanism warrants future well-designed functional experiments to clarify the role of platelets in these cellular processes. In addition, antiplatelet agents, such as purinergic antagonists, are used clinically because they affect inflammatory pathways (48). Recent publications demonstrate that platelets suppress T-cell responses against tumors through production and activation of immunosuppressive factors. These results suggest the use of a combination of immunotherapy and platelet inhibitors, such as aspirin (49, 50) and clopidogrel, as a therapeutic strategy against cancer (51, 52). Therefore, it is possible that antiplatelet therapy could reduce lung cancer risk.

However, we acknowledge some limitations in our study. First, some associations between genetic instrumental variables and phenotype (platelet count) were insufficient and thus may result in a "weak instrument" phenomenon (53). Second, in some scenarios, inconsistent results were observed between INW and MR-Egger (or regular and penalized/robust) models. This phenomenon indicates that genetic variants probably have horizontal pleiotropy, and thus Mendelian randomization assumptions are likely violated (54). Moreover, there is heterogeneity across results incorporating different SNP sets as instrumental variables, which indicates that the instrumental variable should be curated carefully before Mendelian randomization analysis. In this study, all platelet count–related SNPs were curated, and six were retained to better satisfy Mendelian randomization assumptions. Third, a linear association was assumed between platelet count and lung cancer risk. However, the shape could be nonlinear and thus warrants further study incorporating individual-level data. Fourth, we only evaluated platelet count as a potential causal factor, whereas platelet function plays a comparable causal role in this pathway. This may be due to the nature of the data and the lack of detailed platelet information. Future studies should include immature platelet fractions and function analysis. In addition, we assumed that study populations used for the genetic instrument for platelet count and for risk of lung cancer were representative of the same general Caucasian population, which may not be true. Therefore, additional functional studies are needed to further evaluate the mechanisms that underlie associations between platelets and lung cancer risk.

Nonetheless, our findings do suggest a role of platelet count in risk of lung cancer. The results provide a better understanding of lung cancer etiology and evidence for a possible role of antiplatelet interventions in lung cancer prevention.

**Disclosure of Potential Conflicts of Interest**

G. Liu has received speakers bureau honoraria from Pfizer, Astra Zeneca, Takeda, Roche, Novartis, BMS, and Merck. E.H.F.M. van der Heijden reports other commercial research support from Philips Medical Systems and Astra Zeneca Oncology, and has received speakers bureau honoraria from Pentax Medical. No potential conflicts of interest were disclosed by the other authors.

**Disclaimer**

Sponsors had no role in the design of the study, collection and analysis of data, or preparation of the manuscript.

**Authors’ Contributions**

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Analysis and interpretation of data (e.g., statistical analysis, biostatistics, computational analysis): Y. Zhu, Y. Wei, R. Zhang, X. Dong, S. Shen, Y. Zhao, J. Bai, N.E. Caporaso, B. Zhu, P. Brennan, H. Bickeboeller, S.M. Arnold, M.W. Marcus, R.J. Hung, Y. Brarane, L. Kachuri, M.B. Schabath, Y. Han, G.E. Goodman, D.C. Christiani

**Table 2. Association between platelet count and risk of lung cancer using multiple instrumental variable analysis**

<table>
<thead>
<tr>
<th>SNP</th>
<th>Overall NSCLC</th>
<th>Adenocarcinoma</th>
<th>SCC</th>
<th>SCLC</th>
</tr>
</thead>
<tbody>
<tr>
<td></td>
<td>OR (95% CI)</td>
<td>P</td>
<td>OR (95% CI)</td>
<td>P</td>
</tr>
<tr>
<td>IVW</td>
<td>1.62 (1.5–2.7)</td>
<td>0.005</td>
<td>1.51 (0.92–2.48)</td>
<td>0.11</td>
</tr>
<tr>
<td>Penalized IVW</td>
<td>1.62 (1.5–2.7)</td>
<td>0.005</td>
<td>1.51 (0.92–2.48)</td>
<td>0.11</td>
</tr>
<tr>
<td>Robust IVW</td>
<td>1.63 (1.26–2.11)</td>
<td>&lt;0.001</td>
<td>1.51 (0.90–2.53)</td>
<td>0.12</td>
</tr>
<tr>
<td>MR-Egger</td>
<td>3.25 (1.36–9.31)</td>
<td>0.03</td>
<td>6.06 (1.45–25.27)</td>
<td>0.01</td>
</tr>
<tr>
<td>Penalized MR-Egger</td>
<td>3.25 (1.36–9.31)</td>
<td>0.03</td>
<td>6.06 (1.45–25.27)</td>
<td>0.01</td>
</tr>
<tr>
<td>Robust MR-Egger</td>
<td>3.23 (1.80–5.78)</td>
<td>&lt;0.001</td>
<td>5.88 (2.74–12.61)</td>
<td>&lt;0.001</td>
</tr>
</tbody>
</table>

*NOTE: OR of platelet count on lung cancer risk per 100 x 10^3/L increment of platelet count.*
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Other (final acceptance of manuscript): F.A. Shepherd

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