Association of BMI, Smoking, and Alcohol with Multiple Myeloma Mortality in Asians: A Pooled Analysis of More than 800,000 Participants in the Asia Cohort Consortium

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Abstract

Background: To date, few epidemiologic studies have been conducted to elucidate lifestyle-related risk factors for multiple myeloma in Asia. We investigated the association of body mass index (BMI), smoking, and alcohol intake with the risk of multiple myeloma mortality through a pooled analysis of more than 800,000 participants in the Asia Cohort Consortium.

Methods: The analysis included 805,309 participants contributing 10,221,623 person-years of accumulated follow-up across Asia Cohort Consortium cohorts. HRs and 95% confidence intervals (95% CI) for the association between BMI, smoking, and alcohol at baseline and the risk of multiple myeloma mortality were assessed using a Cox proportional hazards model with shared frailty.

Results: We observed a statistically significant dose-dependent association between BMI categories and the risk of multiple myeloma mortality (<18.5 kg/m²: HR = 0.80, 95% CI: 0.52–1.24; 18.5–24.9 kg/m²: reference; 25.0–29.9 kg/m²: HR = 1.17, 95% CI: 0.94–1.47; ≥30 kg/m²: HR = 1.61, 95% CI: 0.99–2.64, P_trend = 0.014). By sex, this association was more apparent in women than in men (P for heterogeneity between sexes = 0.150). We observed no significant associations between smoking or alcohol consumption and risk of multiple myeloma mortality.

Conclusions: This study showed that excess body mass is associated with an increased risk of multiple myeloma mortality among Asian populations. In contrast, our results do not support an association between smoking or alcohol consumption and the risk of multiple myeloma mortality in Asian populations.

Impact: This study provides important evidence on the association of BMI, smoking, and alcohol with the risk of multiple myeloma mortality in Asian populations.
Introduction

Multiple myeloma is the second most common hematologic malignancy and is characterized by the neoplastic proliferation of plasma cells producing monoclonal immunoglobulins (1). Despite recent progress in treatment, multiple myeloma is still an incurable disease associated with substantial mortality. Multiple myeloma occurs in people of all races and from all geographic locations but its incidence varies greatly across regions and countries (2). In general, Western countries show a higher incidence than Asian countries. For example, in the United States, the age-standardized rates (ASR) were 4.3 and 3.0 per 100,000 for males and females, respectively. In contrast, the ASRs in Asia were reported to be 1.0 and 0.7 per 100,000 for males and females, respectively (3). The difference in incidence suggests a substantial difference in risk factor exposure between Western and Asian populations.

To date, several epidemiologic studies have been conducted to elucidate lifestyle-related risk factors for multiple myeloma. Body mass index (BMI), smoking, and alcohol intake are the most intensively examined factors. However, most epidemiologic evidence on these factors has been obtained from studies in Western populations (4–6), whereas evidence from Asian populations is very limited. Because there are large differences in the prevalence of obesity, smoking, and drinking habits (7), as well as of multiple myeloma incidence between the Western and Asian populations (3), large prospective studies of the associations between these factors and multiple myeloma risk in Asian populations are, therefore, needed. Given the limited statistical power due to the small number of multiple myeloma cases in each cohort in Asia, a pooled analysis of the existing cohorts from the Asian Cohort Consortium is one of the ideal approaches to evaluate these relationships.

Here, we investigated the association of body mass index, smoking, and alcohol intake with the risk of multiple myeloma mortality through a pooled analysis of more than 800,000 participants from the Asian Cohort Consortium.

Materials and Methods

Study population

Details of the Asian Cohort Consortium have been described elsewhere (8, 9). Briefly, it is a consortium of cohorts in Asian countries developed to explore the association between genetics, environmental exposure, and the etiology of disease, with sufficient statistical power. Among the cohorts participating in the Asian Cohort Consortium, 16 provided information on multiple myeloma–related deaths during follow-up, as well as data on BMI, smoking, alcohol intake, and potential confounders (sex, age, and education) at baseline. We excluded one cohort with missing data on vital status. Thus, we included 15 cohorts (9 in Japan, 2 in China, 1 in Taiwan, 1 in Korea, 1 in India, and 1 in Singapore) in this pooled analysis. We excluded participants who met any of the following criteria: (i) invalid or missing data on vital status (n = 1,565); (ii) missing data on age or sex (n = 3,163); (iii) invalid or missing data on height or weight (n = 2,290); or (iv) BMI ≤15 kg/m² or >40 kg/m² (n = 16,006). We included a total of 805,309 participants (384,927 men and 420,382 women) in this analysis. The Asian Cohort Consortium coordinating center at the National Cancer Center Japan obtained deidentified individual participant data from all participating cohorts and harmonized it for the statistical analysis.

Pooled analysis of the Asian Cohort Consortium cohorts was approved by the ethical committee of the National Cancer Center Japan (number 2014-041) and each study was approved by respective ethic committees overseeing the participating studies. This analysis was also approved by the Institutional Review Board of Aichi Cancer Center Research Institute.

Exposure data and study outcome

Height and weight at baseline were directly measured in 5 cohorts and obtained via self-report in 10 cohorts. Information on smoking, alcohol intake, and potential confounders was obtained through baseline questionnaires. BMI was calculated as weight [kg]/(height [m])². We categorized BMI according to the guidelines of the World Health Organization (10) as follows: <18.5 (underweight); 18.5–24.9 (normal weight); 25.0–29.9 (overweight); and ≥30 kg/m² (obese). We also applied a five-category BMI classification [<20.0, 20.0–22.4, 22.5–24.9 (reference), 25.0–27.4, and ≥27.5 kg/m²] as used in our previous report of a pooled analysis of BMI and overall mortality in the Asian Cohort Consortium (9). Regarding smoking, participants were classified as never, former, or current smokers, as well as by the cumulative exposure to smoking in pack-years (never smokers; <20 pack-years; and ≥20 pack-years). Alcohol intake was calculated as grams per week to unify data on alcohol intake across the cohorts and then the participants were classified into the following three groups: nondrinkers; intake of 1–149 g/week of ethanol; and intake of ≥150 g/week of ethanol. An exposure period for alcohol intake was the year prior to baseline for SMHS, SWHS, Takayama, KMCC and SCHS, and it was not specified for other cohorts.

Study outcome was defined as death due to multiple myeloma (ICD-9: 203 and ICD-10: C90) during follow-up; cause of death was extracted from death certificates.

Statistical analysis

To determine the relative risk of multiple myeloma mortality associated with BMI, smoking, and alcohol intake, we calculated the HRs and 95% confidence intervals (CI) by a Cox proportional hazards model with shared frailty (STATA command stcox, shared; ref. 11) using pooled individual participant data. An individual cohort was considered as the shared frailty variable to account for between-study heterogeneity. The details of this statistical model were described in previous articles (11, 12). We estimated two types of HR: model 1, which was adjusted for age at baseline (continuous) and sex (men or women); and model 2, which was adjusted for age at baseline, sex, education (none, primary, secondary, trade or technical, university, post university, and missing), and exposures of interest including body mass index (<18.5, 18.5–24.9, 25.0–29.9, ≥30 kg/m²); smoking (never smoker, <20 pack-years, ≥20 pack-years); and alcohol intake (non-drinker, <150 g/week of ethanol, ≥150 g/week of ethanol). Missing values for covariates were treated as dummy variables in the models. We calculated P-values for trend using ordinal variables across each exposure category. We also performed stratified analyses based on geographic regions (East Asia: cohorts from China, Japan, Korea, Singapore, and Taiwan; South Asia: cohorts from India) and countries. Likelihood-based methods were used to test for heterogeneity: between sexes; by smoking status; and across geographic regions and countries. The proportional hazards
assumptions were tested using scaled Schoenfeld residuals and were found to be justified. All statistical analyses were performed using Stata version 14.1 software (Stata Corp.), and \( P < 0.05 \) was considered to be statistically significant.

Results

Table 1 shows the main characteristics of the participating cohorts in the Asian Cohort Consortium. The final analysis included 805,309 participants accounting for 10,221,623 person-years of accumulated follow-up. Mean age at baseline was approximately 54 years in both men and women. Mean BMI ranged from 21.8 to 24.0 kg/m\(^2\) in men and from 21.9 to 24.1 in women. The proportion of ever smokers was 63.7% in men and 6.8% in women, respectively. Mean alcohol intake was 130 g/week in men and 9 g/week in women. During the follow-up period (12.7 years on average), we identified a total of 428 multiple myeloma–related deaths, accounting for 0.33% of deaths from all causes.

Table 2 shows the adjusted HRs for multiple myeloma mortality based on BMI. We observed a statistically significant dose-dependent association between BMI categories and the risk of multiple myeloma mortality (<18.5 kg/m\(^2\); HR = 0.80, 95% CI: 0.52–1.24; 18.5–24.9 kg/m\(^2\); reference; 25.0–29.9 kg/m\(^2\); HR = 1.17, 0.94–1.47; ≥30 kg/m\(^2\); HR = 1.61, 0.99–2.64, \( P_{trend} = 0.014 \), per 1 kg/m\(^2\); HR = 1.04, 1.01–1.07, model 2). By sex, we also observed this significant association among women, but observed no clear association among men (Supplementary Table S1). When we applied a five-category BMI classification, we observed similar results (Table 1; Supplementary Table S1).

Table 3 shows results of the stratified analysis by region and country. Similar findings were observed when the analysis was restricted to East Asia, but evaluation for data from South Asia was difficult because the number of multiple myeloma–related deaths was low. We did not observe significant heterogeneity of association across countries. To evaluate whether smoking modified the association between BMI and the risk of multiple myeloma mortality, we performed stratified analysis by smoking status. The association between BMI and multiple myeloma mortality was more apparent in never-smokers than ever-smokers, although no formal evidence of heterogeneity was observed by smoking status (Supplementary Table S2).

Supplementary Table S3 shows the adjusted HRs for multiple myeloma mortality based on smoking and alcohol intake categories. We observed no association between smoking or alcohol intake and the risk of multiple myeloma mortality. However, these associations were difficult to evaluate in women because the proportions of female smokers and drinkers were low and the number of the female cases who were smokers and drinkers was small.

Finally, we performed several sensitivity analyses as follows: (i) excluding the first 3 and 5 years of follow-up, (ii) excluding participants with a past history of cancer to avoid the possibility of reverse causality, and (iii) excluding female participants in the analysis of smoking and alcohol. These analyses did not change our main results substantially (Supplementary Tables S4–S6). We also conducted another sensitivity analysis by excluding one cohort at a time and ensured that our finding was not driven by any single cohort (Supplementary Table S7).

Discussion

In this pooled analysis of more than 800,000 Asian participants, we found a statistically significant dose-dependent association between BMI and multiple myeloma mortality among Asian populations. By sex, this association this association was more apparent in women than in men. We observed no significant association between smoking or alcohol intake and the risk of multiple myeloma mortality.

A positive association between higher BMI and the risk of multiple myeloma incidence and mortality has been previously reported (4, 12). The 2016 IARC update on body fatness and cancer concluded that excess BMI is a risk factor for multiple myeloma (13). However, although there is sufficient evidence in Western populations, only a few studies have been conducted in Asia. A prospective cohort study, which involved 781,283 Korean men and 103 multiple myeloma cases, did not show a significant association between higher BMI and multiple myeloma incidence, among both sexes combined, with an HR of 0.98 (95% CI: 0.30–3.32) for the BMI category of 27.0–29.9 kg/m\(^2\) relative to the reference category of 18.5–22.9 kg/m\(^2\) (14). The IPHC study in Japan, which is one of the 15 cohort studies in this pooled analysis, also did not report a statistically significant association between BMI and multiple myeloma incidence among both sexes combined (23.0–29.9 kg/m\(^2\); reference: 25.0–29.9 kg/m\(^2\); HR = 0.79, 0.45–1.38; ≥30 kg/m\(^2\); HR = 0.76, 0.45–1.38; ref. 15). Parr and colleagues conducted a pooled analysis of 424,519 participants in the Asia-Pacific Cohort Collaboration and did not observe a statistically significant association between obesity and multiple myeloma mortality for both sexes combined, with an HR of 1.20 (95% CI: 0.59–2.43) for the BMI category of ≥30 kg/m\(^2\) relative to the reference category of 18.5–22.9 kg/m\(^2\) (16). In contrast, the JACC study in Japan, another cohort study included in this pooled analysis, showed a statistically significant association between obesity and multiple myeloma mortality only among women (18.5–25.0 kg/m\(^2\); reference: ≥30 kg/m\(^2\); HR = 4.11, 1.45–11.64; ref. 17). This study observed a statistically significant dose-dependent association between BMI and multiple myeloma mortality also only among women. A recent meta-analysis suggests no sex difference in the association between BMI and multiple myeloma risk in mainly Western populations (4). This discrepancy could be explained by the differences in body fat distribution (18) and metabolic profiles or in genetic susceptibility to obesity (19) between Western and Asian populations. The possible sex difference in Asian populations should be elucidated in future studies.

Different mechanistic pathways for the effect of excess BMI on the development of multiple myeloma have been proposed (20). Adiponectin, an adipocyte-derived cytokine that is inversely correlated with BMI, has been shown to inhibit proliferation of multiple myeloma cells and reduce tumorigenic angiogenesis (21, 22). In support of this hypothesis, Hofmann and colleagues reported that low levels of circulating adiponectin were associated with multiple myeloma risk in overweight and obese individuals (23). They also reported that adiponectin levels were significantly lower among patients with multiple myeloma than among patients with monoclonal gammopathy of undetermined significance (MGUS), the multiple myeloma precursor, suggesting that reduced expression of adiponectin may be associated with progression from MGUS to multiple myeloma (24). A recent analysis showed that there is a large variation in adiponectin levels
### Table 1. Characteristics of the cohort studies in the present pooled analysis

<table>
<thead>
<tr>
<th>Country and study</th>
<th>No. of subjects</th>
<th>Enrollment period</th>
<th>Mean follow-up, years (SD)</th>
<th>Person-years</th>
<th>Mean age at baseline, years (SD)</th>
<th>Mean BMI at baseline (SD)</th>
<th>% of ever smokers</th>
<th>Mean alcohol intake at baseline, g/week (SD)</th>
<th>Method of height and weight ascertained</th>
<th>% of ever smokers at baseline, g/week (SD)</th>
<th>No. of deaths</th>
<th>No. of myeloma deaths</th>
<th>% of myeloma deaths</th>
</tr>
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<td>China</td>
<td></td>
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<td>SMHS</td>
<td>61,426</td>
<td>2001-2006</td>
<td>9.5 (1.8)</td>
<td>581,041</td>
<td>55.4 (9.7)</td>
<td>NA</td>
<td>NA</td>
<td>23.7 (3.1)</td>
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<td>69.6 NA</td>
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<td>5,423</td>
<td>22</td>
</tr>
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<td>14.9 (2.3)</td>
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<td>52.6 (9.1)</td>
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<td>NA</td>
<td>24.0 (3.4)</td>
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<td>NA 2.8 NA</td>
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<tr>
<td>India</td>
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<td>Mumbai</td>
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<td>52.6 (10.9)</td>
<td>48.0 (11.0)</td>
<td>22.1 (3.7)</td>
<td>22.8 (4.5)</td>
<td>DM</td>
<td>31.3 0.4 NA</td>
<td>82 (175)</td>
<td>5,423</td>
<td>22</td>
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<tr>
<td>Japan</td>
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<tr>
<td>3-Pref Aichi</td>
<td>32,142</td>
<td>1984</td>
<td>11.6 (5.0)</td>
<td>340,376</td>
<td>56.5 (11.0)</td>
<td>52.2 (7.4)</td>
<td>23.0 (2.9)</td>
<td>23.3 (3.4)</td>
<td>DM</td>
<td>78.3 8.3 NA</td>
<td>288 (287)</td>
<td>5,465</td>
<td>17</td>
</tr>
<tr>
<td>JPHC1</td>
<td>42,728</td>
<td>1990-1992</td>
<td>21.0 (4.3)</td>
<td>897,432</td>
<td>50.5 (9.5)</td>
<td>49.7 (5.9)</td>
<td>23.6 (2.8)</td>
<td>23.6 (3.1)</td>
<td>DM</td>
<td>75.7 7.5 NA</td>
<td>237 (314)</td>
<td>7,392</td>
<td>37</td>
</tr>
<tr>
<td>JPHC2</td>
<td>55,675</td>
<td>1992-1995</td>
<td>17.7 (4.2)</td>
<td>986,710</td>
<td>54.0 (8.8)</td>
<td>54.4 (8.8)</td>
<td>23.5 (2.9)</td>
<td>23.4 (3.2)</td>
<td>DM</td>
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<td>211 (289) 15 (79)</td>
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<td>3-Pref Miyagi</td>
<td>44,842</td>
<td>1990</td>
<td>16.2 (3.7)</td>
<td>7,258,882</td>
<td>57.6 (10.2)</td>
<td>27.6 (2.8)</td>
<td>22.0 (3.1)</td>
<td>DM</td>
<td>76.2 6.0 NA</td>
<td>NA NA NA</td>
<td>12,851</td>
<td>72</td>
</tr>
<tr>
<td></td>
<td>3-Pref Aichi</td>
<td>44,842</td>
<td>1990</td>
<td>16.2 (3.7)</td>
<td>7,258,882</td>
<td>57.6 (10.2)</td>
<td>27.6 (2.8)</td>
<td>22.0 (3.1)</td>
<td>DM</td>
<td>76.2 6.0 NA</td>
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<td>12,851</td>
<td>72</td>
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<tr>
<td>Miyagi</td>
<td>42,728</td>
<td>1990-1992</td>
<td>21.0 (4.3)</td>
<td>897,432</td>
<td>50.5 (9.5)</td>
<td>49.7 (5.9)</td>
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<td>75.7 7.5 NA</td>
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<tr>
<td>Myagi</td>
<td>42,728</td>
<td>1990-1992</td>
<td>21.0 (4.3)</td>
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<td>75.7 7.5 NA</td>
<td>237 (314)</td>
<td>7,392</td>
<td>37</td>
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<td>Oncraki</td>
<td>47,607</td>
<td>1995</td>
<td>10.8 (4.3)</td>
<td>513,397</td>
<td>59.5 (10.6)</td>
<td>60.7 (9.9)</td>
<td>23.3 (2.9)</td>
<td>23.7 (3.2)</td>
<td>DM</td>
<td>77.5 8.8 18 (16)</td>
<td>187 (186) 16 (68)</td>
<td>7,993</td>
<td>0.49%</td>
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<tr>
<td>RERF</td>
<td>49,424</td>
<td>1963-1995</td>
<td>21.9 (10.3)</td>
<td>1,084,701</td>
<td>52.4 (11.1)</td>
<td>51.9 (10.0)</td>
<td>21.8 (3.0)</td>
<td>21.9 (3.3)</td>
<td>DM</td>
<td>85.4 14.7 204 (262)</td>
<td>15 (63) 25,330</td>
<td>27 (101)</td>
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<td>340,376</td>
<td>56.5 (11.0)</td>
<td>52.2 (7.4)</td>
<td>23.0 (2.9)</td>
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<td>59.3 8.7 NA</td>
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<td>5,848</td>
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<td>Takayama</td>
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<td>1992</td>
<td>13.7 (4.0)</td>
<td>405,102</td>
<td>54.9 (10.2)</td>
<td>55.8 (13.0)</td>
<td>22.5 (2.8)</td>
<td>22.0 (2.9)</td>
<td>DM</td>
<td>81.8 15.9 288 (287)</td>
<td>54 (18) 5,465</td>
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<td>0.31%</td>
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<td>Korea</td>
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<td>KMCC</td>
<td>18,962</td>
<td>1994-2004</td>
<td>13.8 (4.7)</td>
<td>261,165</td>
<td>55.3 (14.5)</td>
<td>53.9 (14.3)</td>
<td>23.1 (3.0)</td>
<td>23.9 (3.4)</td>
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<td>78.3 8.3 12 (40)</td>
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<td>0.23%</td>
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<td>SCHS</td>
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<td>1993-1999</td>
<td>11.5 (3.0)</td>
<td>723,862</td>
<td>56.7 (8.0)</td>
<td>56.3 (8.0)</td>
<td>23.0 (3.1)</td>
<td>23.2 (3.2)</td>
<td>DM</td>
<td>58.0 8.8 25 (80)</td>
<td>3 (17) 10,657</td>
<td>33</td>
<td>0.31%</td>
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<td>CBSCP</td>
<td>23,752</td>
<td>1991-1992</td>
<td>15.2 (2.6)</td>
<td>362,062</td>
<td>48.0 (10.2)</td>
<td>46.6 (9.8)</td>
<td>24.0 (3.2)</td>
<td>24.1 (3.5)</td>
<td>DM</td>
<td>56.2 0.9 NA</td>
<td>NA NA NA</td>
<td>2,755</td>
<td>7</td>
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<td>Total</td>
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</table>

Abbreviations: BMI, body mass index; CBSCP, Community-Based Cancer Screening Project; DM, direct measurement; JACC, The Japan Collaborative Cohort Study; JPHC, Japan Public Health Center-based prospective Study (1 and 2); KMCC, Korea Multi-Center Cancer Cohort; Miyagi, The Miyagi Cohort Study; Mumbai, Mumbai Cohort Study; Oncraki, Oncraki Cohort Study; NA, not available; No., number; 3-Pref Aichi, Three-Prefecture Cohort Study Aichi; 3-Pref Miyagi, Three-Prefecture Cohort Study Miyagi; RERF, Radiation Effects Research Foundation; SCHS, Singapore Chinese Health Study; SA, self-assessment; SMHS, Shanghai Men's Health Study; SWHS, Shanghai Women's Health Study; Takayama, Takayama Cohort Study.
Table 2. Risk of multiple myeloma mortality according to body mass index

<table>
<thead>
<tr>
<th>Body mass index</th>
<th>Person-years</th>
<th>No. of cases</th>
<th>HR (model 1)</th>
<th>P_trend[^a]</th>
<th>Per 1 kg/m²</th>
</tr>
</thead>
<tbody>
<tr>
<td></td>
<td>&lt;18.5</td>
<td>18.5–24.9</td>
<td>25.0–29.9</td>
<td>≥30</td>
<td></td>
</tr>
<tr>
<td>Person-years</td>
<td>619,504</td>
<td>6,961,719</td>
<td>2,356,839</td>
<td>283,561</td>
<td></td>
</tr>
<tr>
<td>No. of cases</td>
<td>22</td>
<td>280</td>
<td>109</td>
<td>17</td>
<td></td>
</tr>
<tr>
<td>HR (model 1)</td>
<td>0.81 (0.52–1.25)</td>
<td>1.00 (Reference)</td>
<td>1.17 (0.93–1.46)</td>
<td>1.60 (0.98–2.61)</td>
<td>0.016</td>
</tr>
<tr>
<td>HR (model 2)</td>
<td>0.80 (0.52–1.24)</td>
<td>1.00 (Reference)</td>
<td>1.17 (0.94–1.47)</td>
<td>1.63 (0.99–2.64)</td>
<td>0.014</td>
</tr>
</tbody>
</table>

NOTE: Model 1, HRs are adjusted for age and sex. Model 2, HRs are adjusted for age, sex, smoking, alcohol intake, and education.

[^a]P_trend values were calculated by assigning scores for categories of body mass index.

Table 3. Risk of multiple myeloma mortality according to body mass index by region and country

<table>
<thead>
<tr>
<th>Body mass index</th>
<th>East Asia</th>
<th>South Asia (India)</th>
<th>China</th>
<th>Japan</th>
<th>Korea</th>
<th>Singapore</th>
<th>Taiwan</th>
</tr>
</thead>
<tbody>
<tr>
<td>Person-years</td>
<td>490,945</td>
<td>128,559</td>
<td>61,477</td>
<td>562,488</td>
<td>11,414</td>
<td>44,396</td>
<td>11,710</td>
</tr>
<tr>
<td>No. of cases</td>
<td>21</td>
<td>1</td>
<td>2</td>
<td>18</td>
<td>0</td>
<td>2</td>
<td>0</td>
</tr>
<tr>
<td>HR (model 1)</td>
<td>0.86 (0.55–1.35)</td>
<td>0.49 (0.06–4.30)</td>
<td>0.70 (0.09–5.14)</td>
<td>0.68 (0.09–5.04)</td>
<td>0.88 (0.54–1.44)</td>
<td>0.85 (0.55–1.45)</td>
<td>0.89 (0.55–1.45)</td>
</tr>
<tr>
<td>HR (model 2)</td>
<td>0.86 (0.55–1.35)</td>
<td>0.54 (0.07–4.63)</td>
<td>1.00 (Reference)</td>
<td>1.00 (Reference)</td>
<td>1.00 (Reference)</td>
<td>1.00 (Reference)</td>
<td>1.00 (Reference)</td>
</tr>
</tbody>
</table>

NOTE: Model 1, HRs are adjusted for age and sex. Model 2, HRs are adjusted for age, sex, smoking, alcohol intake, and education.

[^a]P_trend values were calculated by assigning scores for categories of body mass index, with 1 for <18.5, 2 for 18.5–24.9, 3 for 25.0–29.9, 4 for ≥30 kg/m².

[^b]Heterogeneity for trend among regions or among countries.
Multiple Myeloma Consortium conducted a pooled analysis of nine case–control studies, including 2,670 cases and 11,913 controls, and also did not observe an association with smoking (30). Two studies in Japan were similarly null (31, 32). Consistent with these studies, our results do not support an association between smoking and multiple myeloma risk in Asian populations.

A pooled analysis of 59 case–control studies, including 1,567 cases and 7,296 controls, reported that ever-drinking was associated with reduced risk of multiple myeloma (men: OR = 0.72, 95% CI: 0.59–0.89; women: OR = 0.81, 95% CI: 0.68–0.95; ref. 33). A recent meta-analysis of 26 observational studies reported similar findings [pooled relative risk (RR) = 0.88; 95% CI: 0.79–0.99; ref. 34]. However, most of this evidence has been accumulated in Western populations, whereas evidence in Asian populations remains inconclusive. The JACC study (35) and the PIH study (36) which are participating Japanese cohorts in the Asian Cohort Consortium, did not find an association with alcohol. In the pooled analysis reported here, we found no evidence that alcohol consumption was associated with the risk of multiple myeloma mortality.

This study has several strengths, most importantly the analysis of individual-level data from a large multi-site, multi-country cohort, allowing better detection of possible associations and the calculation of more precise estimates. Furthermore, the prospective design is less susceptible to recall bias than case–control studies. Several limitations also warrant consideration. First, some cohorts collected anthropometric data that were self-reported; although the validity of the self-reported height, weight, and BMI was high among these cohorts (37, 38). Second, as our analyses were conducted using information at a single time point (baseline), we were unable to consider subsequent changes in BMI, smoking habit, and alcohol intake over time. Third, we were unable to consider the effect of other potential confounding factors, including physical activity, MGUS status, and family history of hematologic malignancies. Fourth, the outcome of this analysis was mortality, rather than incidence, and we could not distinguish the possible difference in associations with incidence and survival. However, recent analyses have reported that a higher BMI was associated with longer survival among patients with multiple myeloma (39, 40), suggesting that the strength of the association we report here with higher BMI is more likely to be under- than overestimated. In addition, multiple myeloma was a highly fatal disease during the follow-up period of this study; therefore, the observed associations with multiple myeloma mortality are a fair representation of the association with multiple myeloma incidence. Finally, another potential limitation is the accuracy of diagnosis based on death certificate data. Some fatal multiple myeloma cases might not have been reported correctly due to the lack of diagnostic precision in some areas; nonetheless, it seems probable that any misclassification occurred independently of body size, smoking, and alcohol intake.

In conclusion, this study showed that excess body mass is associated with an increased risk of multiple myeloma mortality among Asian populations. In contrast, our results do not support an association between smoking or alcohol consumption and the risk of multiple myeloma mortality in Asian populations.

Disclosure of Potential Conflicts of Interest
No potential conflicts of interest were disclosed.

Authors’ Contributions

Conception and design: T. Ugai, P. Boffetta, M.S. Pednekar, K.-Y. Yoo, K.S. Chia, M. Inoue, K. Matsuo

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Analysis and interpretation of data (e.g., statistical analysis, biostatistics, computational analysis): T. Ugai, Y.-T. Gao, I. Tsuji, S.K. Park, M.S. Pednekar, Y. Sugawara, K. Wada, C.-J. Chen, M. Inoue, K. Matsuo


Administrative, technical, or material support (i.e., reporting or organizing data, constructing databases): E. Saito, M.S. Rahman, I. Tsuji, C. Nagata, M.S. Pednekar, Y.-B. Xiang, Y. Tomata, M. Inoue, D. Kang, K. Matsuo

Study supervision: H. Ito, P. Boffetta, M.S. Pednekar, K.S. Chia, M. Inoue, K. Matsuo

Other (obtained funding and directed the operation for one of the participating cohorts): X.-O. Shu

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References


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