Energy Intake and Risk of Postmenopausal Breast Cancer: An Expanded Analysis in the Prostate, Lung, Colorectal, and Ovarian Cancer Screening Trial (PLCO) Cohort

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Abstract

Although animal experiments have consistently shown a positive relationship between breast cancer and energy intake, evidence from human studies remains inconclusive. In the Prostate, Lung, Colorectal, and Ovarian Cancer Screening Trial cohort, 29,170 women, aged 55 to 75 years, who successfully completed a food frequency questionnaire (FFQ) at entry (1993-2001), were followed through 2007, and 1,319 incident breast cancers were ascertained (median time from FFQ completion to diagnosis, 4.4 years). Women in the highest quartile of energy intake, relative to the lowest, had modestly, but significantly, increased breast cancer risk [multivariate relative risk (RR), 1.21; 95% confidence interval (95% CI), 1.03-1.42; \( P_{\text{trend}} = 0.03 \)]. The inclusion of body mass index and physical activity in the model reduced risk slightly (RR, 1.18; 95% CI, 1.00-1.39; \( P_{\text{trend}} = 0.07 \)). However, in similar analyses using energy intake from a FFQ administered approximately five years after entry (27,428 women; 866 incident breast cancers; median time from FFQ completion to diagnosis, 2.7 years), women in the highest and lowest quartiles of energy intake had similar risk. When follow-up time after the first FFQ was divided into three 4-year periods, the multivariate RRs for high versus low energy intake increased from 1.21 to 1.37 to 1.55 with increasing time since dietary assessment. Although the divergent results for the two FFQs could be due to subtle questionnaire differences, our findings suggest a modest positive association between energy intake and postmenopausal breast cancer that strengthens with time since dietary assessment.

Introduction

The potential influence of energy intake on risk of breast cancer has been studied in both observational and experimental settings (1). Animal studies have provided consistent evidence that energy restriction reduces breast cancer risk (2-4). Suggested mechanisms for the inhibition of mammary carcinogenesis include changes in cell cycle regulation through decreased cell proliferation and increased apoptosis, reduced oxidative stress and strengthened antioxidant defense, and decreased angiogenesis, which is necessary for tumor growth (1, 3). These mechanisms may be mediated by decreased insulin and insulin-like growth factor I levels as well as increased corticosteroid production (3-6).

In contrast to strong evidence from animal studies, evidence from human populations exposed to famine and energy restriction has been mixed (7-10). Epidemiologic studies investigating the relationship between energy excess and breast cancer risk have also been inconclusive.

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Given the scientific interest in the potential role of energy intake in breast carcinogenesis, the limited number of cohorts that have evaluated this relationship, and our earlier findings, in the current study we reanalyzed the association between energy intake and breast cancer risk in the PLCO cohort, including an additional four years of follow-up and a total of 1,319 incident breast cancer cases. In addition, we incorporated energy intake information from a second, somewhat different, food frequency questionnaire (FFQ) that was completed approximately five years after the initial FFQ.

Materials and Methods

PLCO Study Design. The PLCO, coordinated by the National Cancer Institute in 10 U.S. centers, enrolled from November 1993 to September 2001 approximately 77,000 women and 78,000 men, ages 55 to 74 y, in a randomized, two-arm trial, to determine if screening reduces the incidence and mortality of prostate, lung, colorectal, and ovarian cancer. Female participants randomized to the screening arm received a chest X-ray, flexible sigmoidoscopy and a digital rectal examination, and a CA-125 blood test and transvaginal ultrasound. Those participants randomized to the control arm received care by their usual physicians. A complete description of the PLCO study design has been published (25).

Those enrolled in the trial were asked to complete a baseline questionnaire. This self-administered questionnaire collected information about demographics, medical history, and cancer risk factors, including height and weight at baseline. In addition, those randomized to the screening arm were asked at baseline to complete a FFQ developed for the PLCO, called the PLCO Dietary Questionnaire (DQx).9 The DQx asked about typical frequency of intake during the past year for 137 food items and typical portion size questions for 77 of the food items. The DQx also asked about typical frequency of intake in situ for 114 food items and which assessed diet during the preceding year, was distributed to participants in both the screening and the control arms of the trial beginning in 1998 (26).

Incident breast cancers were ascertained from reports by next of kin and personal physicians, death certificates, and other PLCO study forms. Cases were confirmed using hospital records and pathology reports. As >95% of self-reported breast cancer cases were positively confirmed, self-reported cases were included in our analyses.

Analytic Cohorts. The study population was limited to women randomized to the screening arm of the trial (n = 39,115). Women were excluded from the analysis if they did not complete a baseline questionnaire (n = 951), did not return at least one annual study update questionnaire (n = 387), reported a personal history of breast cancer on the baseline questionnaire (n = 1,334), or reported a personal history of cancer on the baseline questionnaire but did not specify the cancer site (n = 162). Therefore, a total of 36,281 women made up the final study population. These women nearly all self-identified as White, non-Hispanic (88.7%) with the remaining women identifying as Black, non-Hispanic (5.7%), Hispanic (1.6%), Asian (3.3%), Pacific Islander (0.4%), or American Indian (0.3%).

Three analytic cohorts were created from this study population (Table 1). The first was limited to women who completed the DQx (n = 30,074). Further exclusions were made if women were not ages 55 to 75 y at time of entry into the cohort (n = 8), missing quantitative responses to ≥8 food frequency questions (n = 332), or in the highest and lowest 1% for energy intake among all women who completed a baseline questionnaire and a DQx (n = 564); energy intake range of remaining women, 606-3,995 kcal/d. Following these exclusions, 29,170 women (80.4% of study population) remained for the DQx-based analyses. Among these women a total of 1,319 incident breast cancer cases were diagnosed during the follow-up period: 1,106 (83.9%) had been confirmed and 213 (16.1%) were awaiting confirmation. Among confirmed cases, there were 894 invasive (80.8%) and 212 in situ (19.2%) breast cancers.

The second analytic cohort was limited to women in the cohort who completed the DHQ (n = 28,591). Further exclusions were made if women were missing quantitative responses to ≥8 food frequency questions (n = 667) or were in the highest or lowest 1% for energy intake among all female DHQ respondents (n = 496; energy intake range of remaining women, 464-3,663 kcal/d). Thus, 27,428 women (75.6% of study population) were included in the DHQ-based analyses. A total of 806 breast cancer cases were identified during follow-up: 610 cases (75.7%) had been confirmed (80.5% invasive and 19.5% in situ) and 196 (24.3%) were awaiting confirmation.

The third analytic cohort included 24,263 women (66.9% of study population) who completed an acceptable DQx and an acceptable DHQ. A total of 717 women developed breast cancer during the follow-up period.

Statistical Analysis. Entry date into the cohort was defined as the latest of the following dates: randomization, the baseline questionnaire, and the FFQ(s) of interest. Follow-up was terminated at the earliest of the following dates prior to March 31, 2007: for confirmed cases, breast cancer diagnosis date; for self-reported but not yet confirmed cases, the midpoint between the annual study update questionnaire that first reported breast cancer and the annual study update questionnaire that preceded it; and for noncases, including those women lost to follow-up, the date of the most recent annual update questionnaire.

Cox proportional hazards regression modeling was used to estimate adjusted hazard ratios (HR) and 95% CIs of breast cancer incidence by energy intake. Age, in months, was used as the time-dependent variable.

Total energy intake was calculated as the summation across all foods of the frequency of consumption multiplied by both usual portion size (g) and energy content per gram (kcal/g). BMI was calculated by dividing current weight (kg) at baseline by the square of height (m) at baseline.

Analyses utilized three models. The age-adjusted model adjusted for age at entry date. The first multivariate model adjusted for age at entry date, study center (10 U.S. study centers), race (White, non-Hispanic; Black, non-Hispanic; Hispanic; Asian; Pacific Islander; American Indian/Alaskan Native), and accepted breast cancer risk factors: first-degree family history of breast cancer (yes/no), history of benign breast disease (yes/no), age at menarche (≤11, 12-13, 14-15, ≥16 yr), age at first birth (≤19, 20-24, 25-29, ≥30 yr, nulliparous), parity (0, 1, 2, 3, 4, ≥5 live births), age at last menses (<40, 40-44, 45-49, 50-54, ≥55 yr), duration of menopausal hormone therapy use (0, 1-5 y, 6-9 y, ≥10 y), height (continuous), mammography in the last 3 y (0, 1, >1), and education (≤11 y, 12 y or completed high school, post-high school training, college graduate, or postgraduate). A second multivariate model adjusted for the covariates in the first multivariate model as well as for the other components of energy balance (in addition to energy intake): BMI (continuous) and physical activity (0, <1, 1, 2, 3, ≥4 h/wk of vigorous activity). Unless otherwise noted, values from the first multivariate model are presented in the results.

P for trend for energy intake was calculated in two ways using the Wald \( \chi^2 \) test: (a) using the median energy intake of each quartile as a continuous variable and (b) using energy intake as a continuous variable. Unless otherwise noted, P values presented in the results are those from the first method. All correlation coefficients are Spearman's rank order correlations. Statistical analysis was done using SAS 9.1.3.

To examine the possible presence of a time lag effect, we divided follow-up in the DQx and DHQ analytic cohorts into 4-year periods. The first 4-year period of follow-up included all women in the FFQ-specific analytic cohort and truncated follow-up of each woman at four years after her entry date. The second 4-year period started four years after each woman's entry date and truncated follow-up eight years after her entry date. The third 4-year period started eight years after the woman's original entry date and continued until the woman's original exit date. To test for statistical significance of the interaction between energy intake and 4-year follow-up period, we utilized the P value associated with the interaction term in the model. We also did analyses in which the first year of follow-up for each analytic cohort was excluded from the analysis.

In addition, to approximate the previously published study (24), we divided the DQx analytic cohort into earlier and later periods of follow-up. The earlier period of follow-up created a cohort comparable with the previous analysis (29,170 women; 810 incident breast cancers; median follow-up, 5.0 y; median time between DQx and diagnosis, 2.8 y), and the later period of follow-up included only the additional follow-up since the previous study (27,223 women; 519 incident breast cancers; median follow-up, 3.7 y; median time between DQx and diagnosis, 7.2 y).

**Results**

Characteristics of the analytic cohorts in this study are presented in Table 1. With 1,319 incident breast cancer cases, the DQx analytic cohort had 63% more cases than the DHQ analytic cohort and 84% more cases than the combined DQx/DHQ analytic cohort. Median follow-up time was 8.7 years for the DQx analytic cohort but approximately 5.5 years for the DHQ and combined DQx/DHQ analytic cohorts. Median time between FFQ completion and breast cancer diagnosis was 4.4 years in the DQx analytic cohort but only 2.7 years in the DHQ analytic cohort.

Figure 1 shows the relative frequency distributions of daily energy intake for the women who completed both FFQs. The distribution curve for energy intake from the DQx is somewhat wider with a flatter peak than that from the DHQ. Median energy intake was 1,654 kcal/d (interdecile range, 1,041-2,551 kcal/d) based on the DQx, and 1,408 kcal/d (interdecile range, 851-2,217 kcal/d) based on the DHQ. Of the 24,263 women who answered both FFQs, 70% reported a higher energy intake on the DQx. Energy intake estimated by the two instruments was moderately correlated (r = 0.56) among all women completing both FFQs, and among these women stratified by age (r = 0.56 in each 5-year age group). The correlation between the two energy intake estimates did not differ by >0.03 across quartiles of BMI or by levels of physical activity. Sources of calories from the two instruments were similar: the mean percentages of calories from alcohol, carbohydrate, fat, and protein were 2.2%, 55.1%, 27.3%, and 15.5% for the DQx and 2.4%, 52.0%, 30.5%, and 15.1% for the DHQ, respectively.

Energy intake estimated by the DQx was positively, but weakly, associated with both BMI (r = 0.08) and physical activity (r = 0.04). Energy intake estimated by the DHQ was also positively, but even more modestly, associated (r = 0.03 and 0.01 for BMI and physical activity, respectively). The women in the DQx and DHQ analytic cohorts who reported higher energy intake were more likely to be younger, heavier, taller, and more active compared with the women who reported lower energy intake (Table 2). Other breast cancer risk factors were not consistently associated with energy intake.

In the DQx analytic cohort (n = 29,170; 1,319 cases), energy intake was modestly, but statistically significantly, associated with increased risk of breast cancer (Table 3). Breast cancer risk, adjusted for accepted risk factors, was 21% higher (95% CI, 1.03-1.42) among women in the highest quartile of energy intake relative to women in the lowest quartile. The test for trend was also statistically significant (r = 0.03). After adjusting for BMI and physical activity, which may share a causal pathway with energy intake, the RR was reduced slightly to 1.18 (95% CI, 1.00-1.39; P trend = 0.07), but remained statistically significant. RRs were slightly reduced with the exclusion of the 213 self-reported cases (RR comparing highest with lowest quartiles of energy intake, 1.15; 95% CI, 0.96-1.33; P trend = 0.14) and with the exclusion of both the 212 in situ and 213 self-reported cases (RR, 1.16; 95% CI, 0.96-1.41; P trend = 0.04). Risk was also
Energy Intake and Risk of Postmenopausal Breast Cancer

Table 1. Characteristics of the analytic cohorts used in this study

<table>
<thead>
<tr>
<th></th>
<th>DQx</th>
<th>DHQ</th>
<th>Combined DQx/DHQ</th>
</tr>
</thead>
<tbody>
<tr>
<td>Cohort participants</td>
<td>29,170</td>
<td>27,428</td>
<td>24,263</td>
</tr>
<tr>
<td>Incident breast cancer cases</td>
<td>1,319</td>
<td>807</td>
<td>717</td>
</tr>
<tr>
<td>Person-years</td>
<td>245,160</td>
<td>145,474</td>
<td>129,742</td>
</tr>
<tr>
<td>Median age at entry, y</td>
<td>62.0</td>
<td>65.0</td>
<td>65.0</td>
</tr>
<tr>
<td>Median follow-up time (interdecile range), y</td>
<td>8.7 (5.8-11.2)</td>
<td>5.4 (3.0-7.2)</td>
<td>5.6 (3.0-7.2)</td>
</tr>
<tr>
<td>Median time between FFQ completion and breast cancer diagnosis (interdecile range), y</td>
<td>4.4 (1.0-8.6)</td>
<td>2.7 (0.5-5.5)</td>
<td>DQx: 6.0 (3.6-9.4)</td>
</tr>
</tbody>
</table>

The association between energy intake and breast cancer risk was essentially unchanged with the addition of daily alcohol consumption (RR comparing highest with lowest quartiles of energy intake, 1.16; 95% CI, 0.98-1.38; \( P_{\text{trend}} = 0.1 \)).

The different results observed for the DQx and DHQ analytic cohorts could also be explained by the different follow-up times, as the DQx analytic cohort had longer follow-up time and, therefore, longer time elapsed from FFQ completion to breast cancer diagnosis (Table 1). To explore the possible importance of follow-up time, we did analyses on the DQx analytic cohort divided into 4-year periods of follow-up (Table 4). In the DQx analytic cohort, the positive association between risk of breast cancer and energy intake increased with time elapsed since energy intake assessment. Women in the highest versus lowest quartiles of energy intake had a RR of 1.21 (95% CI, 1.00-1.43; \( P_{\text{trend}} = 0.06 \)) in the first four years of follow-up and 1.37 (95% CI, 1.09-1.73; \( P_{\text{trend}} = 0.007 \)) in the second four years of follow-up, and 1.55 (95% CI, 0.89-2.70; \( P_{\text{trend}} = 0.17 \)) in the last years of follow-up. When comparable analyses were done for the DHQ analytic cohort, no association was seen for the first four years of follow-up (RR, 1.00; 95% CI, 0.80-1.24; \( P_{\text{trend}} = 0.95 \)). However, energy intake was weakly, though not statistically significantly, associated with increased breast cancer risk during the remaining years of follow-up (RR, 1.23; 95% CI, 0.77-1.98; \( P_{\text{trend}} = 0.47 \)). The interaction between energy intake and DQx 4-year period of follow-up was of borderline statistical significance (\( P = 0.06 \)); the interaction was not statistically significant in the DHQ analytic cohort (\( P = 0.48 \)).

Figure 1. Frequency distributions of daily energy intake (kcal/d) for women who completed both the DQx and DHQ food frequency questionnaires (\( n = 24,263 \)).


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Table 2. Age-standardized baseline characteristics of the DQx and DHQ analytic cohorts by quartiles of energy intake

<table>
<thead>
<tr>
<th>Range of energy intake (kcal/d)</th>
<th>DQx analytic cohort (n = 29,170)</th>
<th>DHQ analytic cohort (n = 27,428)</th>
</tr>
</thead>
<tbody>
<tr>
<td>Q1</td>
<td>1,095-1,407</td>
<td>1,408-1,796</td>
</tr>
<tr>
<td>Q2</td>
<td>1,312-1,653</td>
<td>1,654-2,080</td>
</tr>
<tr>
<td>Q3</td>
<td>1,095-1,407</td>
<td>1,408-1,796</td>
</tr>
<tr>
<td>Q4</td>
<td>1,312-1,653</td>
<td>1,654-2,080</td>
</tr>
</tbody>
</table>

**NOTE:** All values were standardized to the age distribution (by 5-year categories) of the DQx or DHQ analytic cohorts. Percents may not sum to 100% due to rounding. The DQx analytic cohort included all women in the study population (see Materials and Methods) who completed the DQx food frequency questionnaire (FFQ), were ages 55 to 75 y at entry into the cohort, did not have extreme energy intakes (top or bottom 1% of energy intake among all female DQx respondents). The DHQ analytic cohort included all women in the study population who completed the DHQ FFQ, did not have ≥8 missing responses on the DHQ, and did not have extreme energy intakes (top or bottom 1% of energy intake among all female DHQ respondents).

To determine whether the positive association between energy intake and breast cancer risk in the expanded DQx analytic cohort was attributable to our previously published findings for the same cohort with shorter follow-up time (24), we examined the DQx analytic cohort divided into earlier and later periods of follow-up (see Materials and Methods). In both periods of follow-up, energy intake was positively, but not significantly, associated with breast cancer risk (data not shown). For women in the highest quartile of energy intake relative to those in the lowest quartile, breast cancer risk was increased 15% (95% CI, 0.94-1.41; \( P_{\text{trend}} = 0.18 \)) in the earlier period of follow-up and 25% (95% CI, 0.96-1.63; \( P_{\text{trend}} = 0.13 \)) in the later period of follow-up. The stronger associations during the later period of follow-up indicate that the results of our current analysis extend, and do not simply repeat, our previously published findings (24). In addition, the increased risk lends support to the idea that time elapsed from energy intake assessment to diagnosis may be relevant. However, the interaction between energy intake and period of follow-up was not statistically significant (\( P = 0.70 \)).

In addition, among women who completed both the DQx and the DHQ (n = 24,263; 717 cases), we integrated energy intake from both FFQs in several ways (data not shown). First, we averaged the energy intake percentiles from both FFQs in several ways (data not shown).
with those in the lowest quartile. Second, we averaged the absolute energy intakes of the two FFQs for each woman and compared women in the highest quartile of mean absolute energy intake with those in the lowest. Finally, we ranked women by the quartile of energy intake in each FFQ and compared women jointly in the highest quartile of energy intake for each FFQ with women jointly in the lowest quartile of energy intake for each FFQ. The multivariate RRs of women in the highest level of integrated energy intake, relative to women in the lowest, ranged from 1.16 to 1.22, none of the confidence intervals excluded 1.0, and none of the various tests for trend were statistically significant.

Discussion

In an expanded analysis of a previously published study of energy intake and risk of postmenopausal breast cancer in the PLCO cohort (24), we again observed a modest, but statistically significant, positive association between energy intake and risk of subsequent breast cancer. With four additional years of follow-up time and nearly 75% more breast cancer cases, we found a 20% increase in breast cancer risk comparing extreme quartiles of energy intake. This increase was minimally confounded by BMI and physical activity and strengthened noticeably with time elapsed since energy intake assessment. Results among women in the most recent four years of follow-up seemed stronger than those in the earlier years of follow-up, which indicates that our overall findings were not driven by the results we have already published.

Energy intake estimated from the second FFQ administered in the study, the DHQ, which was given, on average, 3.3 years after the DQx, provided another opportunity to expand upon the previous study. Although energy intake was positively and significantly associated with increased breast cancer risk when using DQx-based energy intake, there was no evidence of an overall association when using DHQ-based energy intake.

The difference in results by FFQ could be explained by the different analytical populations, the different follow-up times, or the different dietary assessment instruments. Firstly, although 1,742 fewer women were included in the DHQ analytic cohort than in the DQx analytic cohort, the two analytic cohorts were very similar, with >24,000 women completing both FFQs and, therefore, in both cohorts. Analyses among the 24,263 women who completed both FFQs showed a positive association between DQx-based energy intake and breast cancer risk comparable with that in the DQx analytic cohort; therefore, the difference in analytic populations is not likely the reason for the difference in results by FFQ.

Secondly, the difference in results by FFQ could be explained by the different follow-up times. Each FFQ quantified energy intake, but the DHQ measured energy intake at a later point in time for each woman: the DHQ was administered, on average, 3.3 years after the DQx. Mean and median follow-up times were 3.1 and 3.3 years longer, respectively, in the DQx analytic cohort than in the DHQ analytic cohort. It is possible that several years need to elapse before the full influence of an increased or decreased energy intake on breast cancer risk becomes apparent. Given more follow-up time, energy intake assessed by the DHQ may also show positive associations with risk, as a statistically significant 38% increase in risk, comparing extreme quartiles, did appear when using DQx-based energy intake among the women completing the DHQ.

Evidence for the importance of elapsed time also came from time lag analyses in the DQx and DHQ analytic cohorts. In 4-year periods of follow-up created from the DQx analytic cohort, the relative risk of breast cancer increased from 1.21 to 1.37 to 1.55 with time elapsed since energy intake assessment. A similar, but weaker, relationship, with relative risk increasing from 1.00 to 1.23 with elapsed time since energy intake assessment, was detected in 4-year periods of follow-up created from the DHQ analytic cohort.

Lastly, differences in the dietary assessment instruments could explain the differential results by FFQ. The

Table 3. Relative risks (RR) and 95% confidence intervals (95% CI) of breast cancer by quartile of energy intake (kcal/d) in the DQx and DHQ analytic cohorts

<table>
<thead>
<tr>
<th>Quartile of energy intake (kcal/d)</th>
<th>Cases</th>
<th>Person-years</th>
<th>Age-adjusted model RR (95% CI)</th>
<th>Multivariate model* RR (95% CI)</th>
</tr>
</thead>
<tbody>
<tr>
<td>DQx (n = 29,170)</td>
<td></td>
<td></td>
<td></td>
<td></td>
</tr>
<tr>
<td>Q1 (&lt;1,311)</td>
<td>281</td>
<td>60,770</td>
<td>1.00</td>
<td>1.00</td>
</tr>
<tr>
<td>Q2 (1,312-1,653)</td>
<td>335</td>
<td>62,267</td>
<td>1.18 (1.01-1.39)</td>
<td>1.12 (0.95-1.32)</td>
</tr>
<tr>
<td>Q3 (1,654-2,080)</td>
<td>339</td>
<td>61,794</td>
<td>1.19 (1.01-1.39)</td>
<td>1.09 (0.93-1.28)</td>
</tr>
<tr>
<td>Q4 (≥2,081)</td>
<td>364</td>
<td>61,329</td>
<td>1.29 (1.10-1.50)</td>
<td>1.21 (1.03-1.42)</td>
</tr>
<tr>
<td>Trend by median of quartile</td>
<td>0.004</td>
<td></td>
<td>0.03</td>
<td>0.03</td>
</tr>
<tr>
<td>Trend by continuous kcal/d</td>
<td>0.003</td>
<td></td>
<td>0.03</td>
<td>0.03</td>
</tr>
<tr>
<td>DHQ (n = 27,428)</td>
<td></td>
<td></td>
<td></td>
<td></td>
</tr>
<tr>
<td>Q1 (&lt;1,095)</td>
<td>191</td>
<td>36,161</td>
<td>1.00</td>
<td>1.00</td>
</tr>
<tr>
<td>Q2 (1,096-1,407)</td>
<td>202</td>
<td>36,658</td>
<td>1.05 (0.86-1.28)</td>
<td>0.98 (0.80-1.21)</td>
</tr>
<tr>
<td>Q3 (1,408-1,797)</td>
<td>210</td>
<td>36,654</td>
<td>1.09 (0.90-1.33)</td>
<td>1.01 (0.82-1.23)</td>
</tr>
<tr>
<td>Q4 (≥1,797)</td>
<td>203</td>
<td>36,029</td>
<td>1.08 (0.88-1.31)</td>
<td>1.05 (0.81-1.32)</td>
</tr>
<tr>
<td>Trend by median of quartile</td>
<td>0.50</td>
<td></td>
<td>0.99</td>
<td></td>
</tr>
<tr>
<td>Trend by continuous kcal/d</td>
<td>0.56</td>
<td></td>
<td>0.93</td>
<td></td>
</tr>
</tbody>
</table>

NOTE: * trend by median of quartile; trend by continuous kcal/d values were calculated using the median energy intake of each quartile as a continuous variable in the model.

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DQx\(^{10}\) is similar to earlier FFQs in its grid format and inclusion of typical foods in the U.S. diet (26). However, it also incorporated the results of extensive cognitive research with volunteers on how to improve dietary assessment methodology (27, 28). Thus, the DQx includes an empirically derived comprehensive food list, rational categories for frequency of intake and portion size, and optimized wording and formatting. The DHQ\(^{11}\) finalized later in time, incorporated the continuing cognitive research (26). It is substantially longer and more comparable to a dietary history with a series of questions after specific food items eliciting additional detail. It does not rely on a conventional grid format and includes somewhat different food items, frequency of intake categories, and portion size definitions than the DQx. The DHQ has been compared to the Block and Willett FFQs in a calibration study using four 24-hour recalls, one in each season, conducted by telephone (29). The DHQ did best overall; with a standard measurement error model, the correlation for energy between estimated truth and the DHQ was only 0.56, and absolute energy intake estimated by the two FFQs only partially overlapped, the correlation between the energy estimates for BMI and 0.01 for physical activity, which suggests that the DQx may be modestly better at estimating energy intake.

In the PLCO cohort, the frequency distributions for energy intake estimated by the two FFQs only partially overlapped, the correlation between the energy estimates was only 0.56, and absolute energy intake estimated by the DHQ was lower than energy intake estimated by the DQx for approximately 70% of the women. It is conceivable, therefore, that the FFQs differ in their ability to assess true energy intake and the DQx is more accurate although the two instruments have not been compared side by side. When we compared sources of calories, the percent of calories from carbohydrate, fat, protein, and alcohol were similar for the two FFQs. Possible explanations for the different performance of the two FFQs include differences in design and respondent burden.

Three other prospective studies have reported positive associations between energy intake and breast cancer risk (17-19). In a large Canadian study of premenopausal and postmenopausal women (with 327,994 and 244,616 person-years of follow-up, respectively), multivariate HRs for highest to lowest quartiles of energy intake, in models including physical activity and BMI, were similar in magnitude to those found in our study and a significant trend was seen (19). After stratification by menopausal status at baseline, the association with energy intake was relatively strong among premenopausal women (multivariate HR,

\(^{10}\) http://prevention.cancer.gov/files/programs-resources/dqx.pdf


Table 4. Relative risks (RR) and 95% confidence intervals (95% CI) of breast cancer by quartile of energy intake (kcal/d) in the DQx and DHQ analytic cohorts divided into 4-year periods of follow-up

<table>
<thead>
<tr>
<th>Quartile of energy intake (kcal/d)</th>
<th>Cases</th>
<th>Person-years</th>
<th>Multivariate model* RR (95% CI)</th>
<th>Quartile of energy intake (kcal/d)</th>
<th>Cases</th>
<th>Person-years</th>
<th>Multivariate model* RR (95% CI)</th>
</tr>
</thead>
<tbody>
<tr>
<td>DQx</td>
<td></td>
<td></td>
<td></td>
<td>DHQ</td>
<td></td>
<td></td>
<td></td>
</tr>
<tr>
<td>First 4 y of follow-up (n = 29,170)</td>
<td></td>
<td></td>
<td></td>
<td></td>
<td></td>
<td></td>
<td></td>
</tr>
<tr>
<td>Q1 (&lt;1,311)</td>
<td>141</td>
<td>28,478</td>
<td>1.00</td>
<td>Q1 (&lt;1,095)</td>
<td>144</td>
<td>25,758</td>
<td>1.00</td>
</tr>
<tr>
<td>Q2 (1,312-1,653)</td>
<td>157</td>
<td>28,467</td>
<td>1.13 (0.96-1.34)</td>
<td>Q2 (1,096-1,407)</td>
<td>143</td>
<td>25,871</td>
<td>0.98 (0.79-1.21)</td>
</tr>
<tr>
<td>Q3 (1,654-2,080)</td>
<td>144</td>
<td>28,585</td>
<td>1.11 (0.94-1.31)</td>
<td>Q3 (1,408-1,796)</td>
<td>152</td>
<td>25,919</td>
<td>1.04 (0.84-1.29)</td>
</tr>
<tr>
<td>Q4 (≥2,081)</td>
<td>157</td>
<td>28,571</td>
<td>1.21 (1.02-1.43)</td>
<td>Q4 (≥1,979)</td>
<td>150</td>
<td>25,785</td>
<td>1.00 (0.80-1.24)</td>
</tr>
<tr>
<td>P trend by median of quartile</td>
<td>0.041</td>
<td></td>
<td></td>
<td>P trend by median of quartile</td>
<td>0.94</td>
<td></td>
<td></td>
</tr>
<tr>
<td>P trend by continuous kcal/d</td>
<td>0.026</td>
<td></td>
<td></td>
<td>P trend by continuous kcal/d</td>
<td>0.97</td>
<td></td>
<td></td>
</tr>
<tr>
<td>Second 4 y of follow-up (n = 27,746)</td>
<td></td>
<td></td>
<td></td>
<td></td>
<td></td>
<td></td>
<td></td>
</tr>
<tr>
<td>Q1 (&lt;1,311)</td>
<td>104</td>
<td>24,278</td>
<td>1.00</td>
<td>Q1 (&lt;1,095)</td>
<td>47</td>
<td>10,404</td>
<td>1.00</td>
</tr>
<tr>
<td>Q2 (1,312-1,653)</td>
<td>135</td>
<td>24,443</td>
<td>1.14 (0.90-1.46)</td>
<td>Q2 (1,096-1,407)</td>
<td>59</td>
<td>10,786</td>
<td>1.19 (0.75-1.90)</td>
</tr>
<tr>
<td>Q3 (1,654-2,080)</td>
<td>132</td>
<td>24,677</td>
<td>1.17 (0.92-1.49)</td>
<td>Q3 (1,408-1,796)</td>
<td>58</td>
<td>10,734</td>
<td>1.39 (0.88-2.19)</td>
</tr>
<tr>
<td>Q4 (≥2,081)</td>
<td>146</td>
<td>24,545</td>
<td>1.37 (1.09-1.73)</td>
<td>Q4 (≥1,979)</td>
<td>53</td>
<td>10,244</td>
<td>1.23 (0.77-1.98)</td>
</tr>
<tr>
<td>P trend by median of quartile</td>
<td>0.007</td>
<td></td>
<td></td>
<td>P trend by median of quartile</td>
<td>0.007</td>
<td></td>
<td></td>
</tr>
<tr>
<td>P trend by continuous kcal/d</td>
<td>0.009</td>
<td></td>
<td></td>
<td>P trend by continuous kcal/d</td>
<td>0.43</td>
<td></td>
<td></td>
</tr>
<tr>
<td>Last years of follow-up† (n = 15,864)</td>
<td></td>
<td></td>
<td></td>
<td></td>
<td></td>
<td></td>
<td></td>
</tr>
<tr>
<td>Q1 (&lt;1,311)</td>
<td>36</td>
<td>8,014</td>
<td>1.00</td>
<td>Q1 (&lt;1,095)</td>
<td>47</td>
<td>10,404</td>
<td>1.00</td>
</tr>
<tr>
<td>Q2 (1,312-1,653)</td>
<td>43</td>
<td>8,358</td>
<td>1.36 (0.77-2.37)</td>
<td>Q2 (1,096-1,407)</td>
<td>59</td>
<td>10,786</td>
<td>1.19 (0.75-1.90)</td>
</tr>
<tr>
<td>Q3 (1,654-2,080)</td>
<td>63</td>
<td>8,532</td>
<td>1.80 (1.06-3.07)</td>
<td>Q3 (1,408-1,796)</td>
<td>58</td>
<td>10,734</td>
<td>1.39 (0.88-2.19)</td>
</tr>
<tr>
<td>Q4 (≥2,081)</td>
<td>61</td>
<td>8,212</td>
<td>1.55 (0.89-2.70)</td>
<td>Q4 (≥1,979)</td>
<td>53</td>
<td>10,244</td>
<td>1.23 (0.77-1.98)</td>
</tr>
<tr>
<td>P trend by median of quartile</td>
<td>0.17</td>
<td></td>
<td></td>
<td>P trend by median of quartile</td>
<td>0.47</td>
<td></td>
<td></td>
</tr>
<tr>
<td>P trend by continuous kcal/d</td>
<td>0.21</td>
<td></td>
<td></td>
<td>P trend by continuous kcal/d</td>
<td>0.43</td>
<td></td>
<td></td>
</tr>
</tbody>
</table>

NOTE: P trend by median of quartile values were calculated using the median energy intake of each quartile as a continuous variable in the model.

P trend by continuous kcal/d Values were calculated using kcal/d as a continuous variable in the model.

*Adjusted for age at entry date, study center (10 PLCO study centers), race (White, non-Hispanic; Black, non-Hispanic; Hispanic; Asian; Pacific Islander; American Indian/Alaskan Native), family history of breast cancer (yes/no), history of benign breast disease (yes/no), age at menarche (≤11, 12-13, 14-15, ≥16 y), age at first birth (≤19, 20-24, 25-29, ≥30 y, nulliparous), number of live births (0, 1, 2, 3, 4, ≥5), age at last menses (<40, 40-44, 45-49, 50-54, ≥55 y), duration of menopausal hormone therapy use (0, 1-5 y, 6-9 y, ≥10 y), and education (≤11 y, 12 y or completed high school, post-high school training, college graduate, or postgraduate), height (continuous), and mammography in the last three years (0, 1, >1 times).

†Follow-up for this period began 4 y after the entry date for each woman and continued through her final exit date from the cohort.

‡Follow-up for this period began 8 y after the entry date for each woman and continued through her final exit date from the cohort.
Energy Intake and Risk of Postmenopausal Breast Cancer

1.45; 95% CI, 1.13-1.85; P\text{trend} = 0.001) but null among postmenopausal women (multivariate HR, 0.94; 95% CI, 0.72-1.23; P\text{trend} = 0.86) with a marginally significant P value for interaction by menopausal status (P = 0.06; ref. 19). Additionally, a small U.S. study of 590 postmenopausal women found increased breast cancer risk with each 500-kilocalorie increase in total energy intake (RR, 2.72; 95% CI, 1.51-4.89; ref. 17), and a Norwegian study of premenopausal and postmenopausal women found evidence of an energy intake-breast cancer association (RR, 1.50; 95% CI, 1.05-2.15) although the association lost statistical significance when BMI was added to the model (18). Case-control studies in Argentina, China, Italy, and Switzerland have provided further evidence supporting a positive association between energy intake and risk of breast cancer (12-16), as have international correlation studies of total calories and breast cancer incidence (r = 0.70) and breast cancer mortality (r = 0.60; ref. 30).

Surprisingly few epidemiologic studies have assessed the relationship between total energy consumption and breast cancer risk. The vast majority of studies of diet and breast cancer adjust nutrient estimates for total calorie intake by one of several accepted methods. Justifications include correcting for systematic underestimation and overestimation of consumption, reducing measurement error by controlling for variation in energy intake, controlling for confounding by total energy intake, and focusing on dietary composition rather than absolute nutrient intake (31). Therefore, most studies, by adjusting for energy, have not been able to evaluate its independent role.

However, our findings do differ from those of three recent cohort studies and one recent case-control study (20-23). Energy intake was not associated with risk among postmenopausal women in the Breast Cancer Detection Demonstration Project Follow-up Cohort Study (20) nor among premenopausal and postmenopausal women in the Nurses’ Health Study cohort (21); both studies presented results controlled for BMI but not physical activity. There was an association among premenopausal and postmenopausal women in the California Teachers Study cohort (22) nor among premenopausal and postmenopausal women in the case-control Shanghai Breast Cancer Study, which did not adjust for BMI or physical activity (23).

The evidence from studies of populations exposed to severe energy restriction has been mixed (7-10). A retrospective cohort study in Sweden found that in a group of women hospitalized for anorexia nervosa before age 40, calorie restriction was associated with a statistically significant 53% decreased risk of breast cancer, compared with the general female Swedish population (7). In a Norwegian study, those women who had undergone puberty during World War II experienced lower breast cancer incidence than those who were older or younger during the same time period. As daily energy intake was 22% lower during the war, caloric restriction may have contributed to the lower incidence (8). In contrast, a case-cohort study of women exposed to the briefer, but more severe famine during World War II reported a significant 53% decreased risk of breast cancer, compared with those in the lowest. When follow-up time was divided into 4-year periods, the RRs for high versus low energy intake increased from 1.21 to 1.37 to 1.55, with those in the highest quartile of energy intake compared with those in the lowest. When follow-up time was divided into 4-year periods, the RRs for high versus low energy intake increased from 1.21 to 1.37 to 1.55, with those in the lowest. When follow-up time was divided into 4-year periods, the RRs for high versus low energy intake increased from 1.21 to 1.37 to 1.55, with those in the lowest. When follow-up time was divided into 4-year periods, the RRs for high versus low energy intake increased from 1.21 to 1.37 to 1.55, with those in the lowest.

Table 3 shows the results of the diet and breast cancer association, including adjusted and not adjusted for BMI or physical activity. There were no significant associations between diet and breast cancer that strengthened after adjustment. However, the relationship between total energy consumption and breast cancer remains significant, as has been reported in other studies (19).

In conclusion, the findings from this extended analysis continue to suggest a modest positive association between total energy intake and risk of postmenopausal breast cancer, an association which was largely independent of BMI and physical activity. In this large cohort, breast cancer risk increased 21% (95% CI, 3-42%) among women in the highest quartile of energy intake compared with those in the lowest. When follow-up time was divided into 4-year periods, the RRs for high versus low energy intake increased from 1.21 to 1.37 to 1.55, with increasing time since dietary assessment. Our findings suggest a modest positive association between energy intake and postmenopausal breast cancer that strengthens with time since dietary assessment.

Disclosure of Potential Conflicts of Interest
No potential conflicts of interest were disclosed.

References
## Energy Intake and Risk of Postmenopausal Breast Cancer: An Expanded Analysis in the Prostate, Lung, Colorectal, and Ovarian Cancer Screening Trial (PLCO) Cohort

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