Investigation of Dietary Factors and Endometrial Cancer Risk Using a Nutrient-wide Association Study Approach in the EPIC and Nurses’ Health Study (NHS) and NHSII


Abstract

Data on the role of dietary factors in endometrial cancer development are limited and inconsistent. We applied a ‘nutrient-wide association study’ approach to systematically evaluate dietary risk associations for endometrial cancer while controlling for multiple hypothesis tests using the false discovery rate (FDR) and validating the results in an independent cohort. We evaluated endometrial cancer risk associations for dietary intake of 84 foods and nutrients based on dietary questionnaires in three prospective studies, the European Prospective Investigation into Cancer and Nutrition (EPIC; N = 1,303 cases) followed by validation of nine foods/nutrients (FDR < 0.10) in the Nurses’ Health Studies (NHS/NHSII; N = 1,531 cases). Cox regression models were used to estimate HRs and 95% confidence intervals (CI). In multivariate adjusted comparisons of the extreme categories of intake at baseline, coffee was inversely associated with endometrial cancer risk (EPIC, median intake 750 g/day vs. 8.6; HR, 0.81; 95% CI, 0.68–0.97, P_trend = 0.09; NHS/NHSII, median intake 1067 g/day vs. none; HR, 0.82; 95% CI, 0.70–0.96, P_trend = 0.04). Eight other dietary factors that were associated with endometrial cancer risk in the EPIC study (total fat, monounsaturated fat, carbohydrates, phosphorus, butter, yogurt, cheese, and potatoes) were not confirmed in the NHS/NHSII. Our findings suggest that coffee intake may be inversely associated with endometrial cancer risk. Further data are needed to confirm these findings and to examine the mechanisms linking coffee intake to endometrial cancer risk to develop improved prevention strategies.

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Nutrient-wide Association Study of Endometrial Cancer

Introduction

Higher endometrial cancer incidence rates in North America and Europe versus lower rates in Africa and South Asia (1) may be explained by endometrial cancer risk factors such as estrogen exposure and obesity, which are related to a westernized lifestyle (2); as diet is an important component of the westernized lifestyle, we hypothesized that dietary factors may also contribute to endometrial cancer etiology. The recent World Cancer Research Fund (WCRF) meta-analysis of 3,571 endometrial cancer cases reported an inverse association between coffee drinking and endometrial cancer risk (3); however, an important unanswered question is whether other dietary factors may play a role in endometrial cancer development.

We sought to investigate whether dietary factors were related to endometrial cancer risk by applying a ‘nutrient-wide association study’ (NWAS) approach (4–6). The NWAS is an application of methods developed for GWAS to identify associations between dietary intake and risk of disease and includes adjustment for multiple comparisons by calculating the false discovery rate (FDR) (7) followed by external validation of results in an independent study. This method has been used to identify novel dietary risk associations for diabetes and blood pressure (4, 6). The current study is the first to use the NWAS method to prospectively evaluate dietary factors and risk of endometrial cancer in European and North American populations.

Materials and Methods

This NWAS investigated intakes of 84 foods/nutrients in relation to endometrial cancer risk in the European Prospective Investigation into Cancer and Nutrition (EPIC) study, calculated the associated FDR to select dietary factors and evaluated these factors and endometrial cancer risk in the validation cohorts, the Nurses’ Health Study (NHS) and NHSII (Supplementary Fig. S1).

Study populations

The EPIC study includes 521,330 participants 25 to 70 years at enrollment (1992–2000; ref. 8). From 367,903 women in the EPIC study, individuals were excluded if they: reported a prevalent cancer except non-melanoma skin cancer or those who died (NHS ¼ 3,630; NHSII ¼ 1,321); those who reported a hysterectomy (NHS ¼ 20,657; NHSII ¼ 6,424); or were ineligible (e.g., duplicate ID; NHS ¼ 36; NHSII ¼ 199). Participants did not contribute person–time in cycles in which they were missing body mass index (BMI) or if they had a total caloric intake (<600 or >3,500 kcal/day) for the expanded food frequency questionnaire (FFQ; 1984 and thereafter), but they could re-enter the analysis for subsequent periods after these data became available. Informed consent was provided by all participants and the study design, data collection, and analyses were performed in accordance with the ethical standards of the institutional review board at the Brigham and Women’s Hospital (Boston, MA).

Ascertainment of endometrial cancer cases

In the EPIC study, incident endometrial cancers were identified through population-based cancer registries or active follow-up, and mortality data were obtained from cancer or mortality registries (8). Tumors were classified as ICD-10 code C54. In total, 1,504 cases were identified and cases were censored if they were not the first incident tumor (n ¼ 54), noninvasive (n ¼ 68), or missing tumor behavior (n ¼ 79). Analyses of type I/II tumors were not conducted in the EPIC study because data on tumor histologic subtype and grade were incomplete.

In the NHS/NHSII, information about new diagnoses was collected in each questionnaire. When a woman reported a cancer, permission was sought to obtain the relevant medical records and pathology reports, and study physicians reviewed these documents to confirm the diagnosis. Deaths in the cohort were classified in the top or bottom 1% of energy intake to energy requirement (n ¼ 6,045); were missing a lifestyle questionnaire (n ¼ 22); or had outlying values for specific nutrient intakes (n ¼ 7), leaving 301,107 participants in this study. Informed consent was provided by all participants and ethical approval for the study was obtained from the internal review board of the International Agency for Research on Cancer and from local ethics committees in each participating country.

The NHS was established in 1976 among 121,700 married, female registered nurses, ages 30 to 55 years, while the NHSII began in 1989 among 116,430 female registered nurses, ages 25 to 42 years (9, 10). Participants excluded at the study baseline (1980 for NHS and 1991 for NHSII) were women with a diagnosis of cancer except non-melanoma skin cancer or those who died (NHS ¼ 3,630; NHSII ¼ 1,321), those who reported a hysterectomy (NHS ¼ 20,657; NHSII ¼ 6,424); or were ineligible (e.g., duplicate ID; NHS ¼ 36; NHSII ¼ 199). Participants did not contribute person–time in cycles in which they were missing body mass index (BMI) or if they had a total caloric intake (<600 or >3,500 kcal/day) for the expanded food frequency questionnaire (FFQ; 1984 and thereafter), but they could re-enter the analysis for subsequent periods after these data became available. Informed consent was provided by all participants and the study design, data collection, and analyses were performed in accordance with the ethical standards of the institutional review board at the Brigham and Women’s Hospital (Boston, MA).

Note: Supplementary data for this article are available at Cancer Epidemiology, Biomarkers & Prevention Online (http://cebp.aacrjournals.org/).

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were identified by reports from family members, the U.S. Postal Service and the National Death Index. Cases were confirmed as epithelial endometrial cancer (NHS = 1,254; NHSII = 277) and of these we further subclassified endometrial cancer as invasive (≥stage II) endometrial adenocarcinoma (NHS = 753; NHSII = 146).

**Dietary assessment**

The diet of the EPIC participants was assessed using validated dietary questionnaires or food records (8) and this analysis evaluated foods that were available in ≥8 countries (Supplementary Materials and Methods). The EPIC Nutrient Database was used to calculate standardized nutrient intake for the 10 countries and all standardized priority nutrients were analyzed.

In the NHS/NHSII, intakes of selected foods/nutrients were assessed from 1980 FFQ (NHS) or 1991 FFQ (NHSII) and every approximately 4 years thereafter until the end of follow-up using a validated and reproducible FFQ (11). Nutrient intakes were calculated by multiplying the frequency of intake by the nutrient content of specified portions based on the U.S. Department of Agriculture food composition data.

**Measurement of other covariates**

The following risk factors for endometrial cancer were adjusted for in all multivariate models: BMI, total energy, smoking status, age at menarche, oral contraceptive (OC) use, parity, and a combined variable for menopausal status and postmenopausal hormone (PMH) use (Supplementary Materials and Methods). In the EPIC study, we additionally adjusted the multivariate models for physical activity, height, education level, and alcohol intake and the risk estimates were very similar; therefore, these covariates were not included in the final models. Family history of endometrial cancer was not available.

**Statistical analysis**

Cox proportional hazards (PH) regression was used to estimate the HRs and 95% confidence intervals (CI). In the EPIC study, age was the underlying time metric for Cox regression with the subjects’ age at recruitment as the entry time and their age at cancer diagnosis, death, emigration, or last follow-up as the exit time. Nutrient intakes were energy-adjusted using the regression residual method and levels of food/nutrient intake were categorized into quartiles unless stated otherwise. To account for multiple comparisons, we estimated the FDR for each food/nutrient, which is the ratio of the number of false positives to the total number of positive associations, or the percentage of findings drawn from the null distribution at a given significance level (7). To compute the FDR, we used an analytic method that estimates the number of false positive results by creating a “null distribution” of regression test statistics; this was accomplished by randomly assigning the endometrial cancer case status, running the Cox proportional hazards model and collecting the associated P value over 1,000 permutations (4, 6).

In the NHS/NHSII, we evaluated dietary intake at the study baseline as well as the cumulative average intake for foods/nutrients that were identified in the analysis in the EPIC study with FDR≤0.10 (Supplementary Material and Methods). The cumulative average intake analyses included a 2- to 6-year time lag between the diet assessment and the start of follow-up; thus each participant accrued person-time beginning with the 1984 (NHS) or 1995 (NHSII) FFQ until their date of endometrial cancer or other cancer diagnosis, hysterectomy, death or the end of follow-up (NHS: June 1, 2010; NHSII: June 1, 2011). Cox PH regression was carried out using age (months) and the biannual questionnaire cycle as the time scale.

The P value for the test of linear trend was calculated by assigning participants the median value for each dietary intake category and this variable was modeled as a continuous term. Random effects meta-analysis was used to combine HRs across studies. Analyses were performed using ‘survival’ and ‘rmeta’ packages in R v3.0.2.

**Results**

In a total of 2,834 incident endometrial cancers that were evaluated, 1,303 were from the EPIC study (mean follow-up = 11 years) and 1,531 were from the NHS/NHSII (mean follow-up = 25 years). The characteristics of the study populations are summarized in Table 1. Of the 84 foods/nutrients that were evaluated in the EPIC study, 10 were associated with endometrial cancer risk (FDR≤0.10) including butter, yogurt, cheese, potatoes, coffee, cream desserts, total fat, monounsaturated fat, carbohydrates, and phosphorus (Supplementary Table S1) while the remainder was not included in the analysis in the NHS/NHSII.

**Table 1.** Age-standardized characteristics at baseline of the EPIC study and the NHS and NHSII

<table>
<thead>
<tr>
<th></th>
<th>EPIC^a</th>
<th>NHS^b</th>
<th>NHSII^c</th>
</tr>
</thead>
<tbody>
<tr>
<td>Participants, n</td>
<td>301,367</td>
<td>68,063</td>
<td>87,543</td>
</tr>
<tr>
<td>Means (SD)</td>
<td>49.8 (9.9)</td>
<td>45.3 (7.2)</td>
<td>35.9 (4.6)</td>
</tr>
<tr>
<td>Age, y</td>
<td>13.1 (1.5)</td>
<td>12.4 (1.8)</td>
<td>12.4 (1.4)</td>
</tr>
<tr>
<td>Body mass index, kg/m²</td>
<td>24.9 (4.4)</td>
<td>24.3 (4.5)</td>
<td>24.5 (5.3)</td>
</tr>
<tr>
<td>Total energy (kcal/day)</td>
<td>1934 (541)</td>
<td>1570 (501)</td>
<td>1794 (548)</td>
</tr>
<tr>
<td>Percentages</td>
<td></td>
<td></td>
<td></td>
</tr>
<tr>
<td>Parous</td>
<td>84</td>
<td>94</td>
<td>73</td>
</tr>
<tr>
<td>Ever use oral contraceptives</td>
<td>60</td>
<td>50</td>
<td>83</td>
</tr>
<tr>
<td>Postmenopausal</td>
<td>44</td>
<td>27</td>
<td>1</td>
</tr>
<tr>
<td>Ever smoked</td>
<td>47</td>
<td>57</td>
<td>1</td>
</tr>
<tr>
<td>Ever use postmenopausal hormones</td>
<td>18</td>
<td>8</td>
<td>1</td>
</tr>
</tbody>
</table>

^aExposures at the study baseline (1980 in the NHS; 1991 in the NHSII).

^bValues are not age-standardized.
Nutrient-wide Association Study of Endometrial Cancer

Discussion

We used the NWAS method to examine consumption of 84 foods/nutrients in the EPIC study and identified 10 dietary factors for which the highest versus lowest consumption levels were associated with increased risk (butter, yogurt, potatoes, carbohydrates) or decreased risk (cheese, coffee, cream desserts, total fat, monounsaturated fat, phosphorus) of endometrial cancer (FDR < 0.10). The inverse association between coffee intake and endometrial cancer risk was confirmed in the NHS/NHSII, suggesting that the other associations in our discovery effort may be false positives. Butter intake was positively associated with endometrial cancer risk in the meta-analysis of all three studies; to our knowledge, butter has not been investigated in previous prospective studies.

The inverse association between coffee intake and endometrial cancer risk is consistent with two recent meta-analyses (3, 12) and an earlier NHS report (13). Obesity is a strong risk factor for endometrial cancer (2) and may act by increasing exposure to estrogen (14) and/or hyperinsulinemia (15). Coffee may relate to these exposures; for example, in a cross-sectional study of >2,000 healthy NHS women, a high coffee intake was associated with lower levels of C-peptide which suggests a possible reduction in insulin secretion by coffee drinkers (16). Increasing consumption of caffeine and caffeine-containing coffee was associated with higher levels of adiponectin in the NHS which may benefit insulin sensitivity (17), and also with elevated levels of sex hormone-binding globulin (SHBG) that may decrease bioavailable estrogen (3) that the decreased risk for endometrial cancer is similar in magnitude for decaffeinated coffee, and because drinking tea which also contains caffeine does not appear to be related to endometrial cancer risk. It is possible that other participant characteristics including differences in dietary intake levels for certain foods/nutrients, could explain why some of the dietary associations were not confirmed. Of those dietary factors that were investigated but not confirmed, consumption of carbohydrates (19–21) and total fat (22) were previously evaluated and, consistent with our final conclusions, there was no association with endometrial cancer risk. It is possible that other participant characteristics that differed between the EPIC, NHS, and NHSII cohorts, including differences in dietary intake levels for certain foods/nutrients, could explain why some of the dietary associations were not confirmed. A meta-analysis of a large number of cohort studies would therefore be useful to further evaluate a range of nutrients/foods in relation to endometrial cancer risk.

Figure 1.
“Manhattan plot” showing results from the nutrient-wide association study method to evaluate the association between dietary intake of various foods and nutrients and endometrial cancer risk in the European Prospective Investigation into Cancer and Nutrition (EPIC) study. The y-axis indicates the \( -\log_{10} \) of the FDR \( P \) value of the multivariate adjusted Cox proportional hazards regression coefficient for the comparison of extreme quartiles or categories of dietary intake. Colors represent the dietary intake categories that were evaluated. Within each category, factors are arranged from left to right in order from the lowest to highest HR. The red horizontal line is \( -\log_{10}(0.10) \). Dietary factors that were selected for confirmation in the NHS/NHSII are labeled with the HR from the EPIC study for the comparison of the highest versus lowest category of dietary intake in relation to risk of endometrial cancer.

respectively, had higher total energy intake (mean = 1,934 kcal/day vs. 1,794) and consumed more yogurt (mean = 65.1 g/day vs. 32.6), cheese (mean = 39.2 g/day vs. 13.6), and butter (mean = 4.0 g/day vs. 2.6) (Table 1 and Supplementary Table S3). Cream dessert was not available in the NHS/NHSII cohorts. Of the nine foods/nutrients that were investigated in the NHS/NHSII, only the association with coffee was replicated (highest vs. lowest categories of intake; overall HR, 0.82; 95% CI, 0.73–0.92; Fig. 2; Supplementary Table S4). A positive association with butter intake was observed in the meta-analysis of all three cohorts (highest versus lowest categories of intake; overall HR, 1.14; 95% CI, 1.02–1.27).

In sensitivity analyses in the NHS/NHSII, there were similar associations for the cumulative average diet (Supplementary Table S5) and when restricting the case group to invasive endometrial adenocarcinomas (Supplementary Table S6). There were no consistent differences in the risk associations when stratifying by BMI or PMH use in the EPIC study or the NHS/NHSII (data not shown).

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The objective of the current study was to use the NWAS method as it may identify novel dietary risk associations with disease as demonstrated by studies of blood pressure and diabetes (4, 6). Advantages of this study design were the ability to systematically demonstrate by studies of blood pressure and diabetes (4, 6).

Figure 2. Forest plots showing multivariate HRs and 95% CIs for comparisons of the highest versus lowest categories of intake of nine selected foods and nutrients reported in the baseline dietary assessment in relation to endometrial cancer risk in the European Prospective Investigation into Cancer and Nutrition (EPIC) and the Nurses’ Health Study (NHS/NHSII). Foods and nutrients were evaluated if they had a FDR < 0.05 for the comparison of extreme quartiles or categories of dietary intake in the EPIC study. The overall multivariate adjusted HR (95% CI) was estimated using random effects meta-analysis. *P values for heterogeneity (Phet) comparing the EPIC study and the NHS/NHSII and the following exceptions: phosphorus (Phet = 0.03) and potatoes (Phet = 0.05). Multivariate models were adjusted for BMI, total energy intake, smoking status, age at menarche, oral contraceptive use, parity, and a combined variable for menopausal status and postmenopausal hormone use and were stratified by age and study center (EPIC) or stratified by age, cohort, and the 2-year questionnaire cycle (NHS/NHSII). Contrasts and median intake values were total fat [EPIC, quartile 4 (Q4, 73.7 g/day) vs. quartile 1 (Q1, 49.7 g/day); NHS/NHSII, Q4, 81.0 g/day vs. Q1, 52.0 g/day], monounsaturated fat (EPIC, Q4, 29.0 g/day vs. Q1, 16.2 g/day; NHS/NHSII, Q4, 34.0 g/day vs. Q1, 19.1 g/day), carbohydrates (EPIC, Q4, 212.0 g/day vs. Q1, 148.1 g/day; NHS/NHSII, Q4, 247.0 g/day vs. Q1, 131.0 g/day), phosphorus (EPIC, Q4, 1490 mg/day vs. Q1, 984 mg/day; NHS/NHSII, Q4, 1579 mg/day vs. Q1, 951 mg/day), butter (EPIC, highest, 10.3 g/day vs. lowest, 0 g/day; NHS/NHSII, highest, 5.0 g/day vs. lowest, 0 g/day), yogurt (EPIC, Q4, 143.5 g/day vs. Q1, 0 g/day; NHS/NHSII, highest, 105.4 g/day vs. lowest, 0 g/day), cheese (EPIC, Q4, 74.7 g/day vs. Q1, 7.5 g/day; NHS/NHSII, highest, 28.0 g/day vs. lowest, 2.0 g/day), potatoes (EPIC, Q4, 142.2 g/day vs. Q1, 20.8 g/day; NHS/NHSII, highest, 146.2 g/day vs. lowest, 25.9 g/day), and coffee (EPIC, Q4, 750.0 g/day vs. Q1, 8.6 g/day; NHS/NHSII, highest, 1066.5 g/day vs. lowest, 0 g/day).

The authors assume full responsibility for analyses and interpretation of these data.

Authors’ Contributions


No potential conflicts of interest were disclosed.

Disclosure of Potential Conflicts of Interest

No potential conflicts of interest were disclosed.

Disclaimer

The authors assume full responsibility for analyses and interpretation of these data.

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