A Systematic Review of p53 as a Prognostic Factor of Survival in Squamous Cell Carcinoma of the Four Main Anatomical Subsites of the Head and Neck

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Abstract

Objectives: To summarize existing evidence about whether the presence of mutant or upregulated p53 is a prognostic factor for patients presenting with squamous cell carcinoma arising from the larynx, oropharynx, hypopharynx, or oral cavity.

Method: Relevant articles were identified using strict criteria for systematic searches. Associations between mutant or upregulated p53 versus wild-type or low/undetectable p53 in relation to overall survival and DFS were summarized by extracting or deriving hazard ratio (HR) estimates. Random-effects meta-analyses were used to account for between-study heterogeneity and to summarize the effect of p53 across studies.

Results: The meta-analyses gave a statistically significant pooled HR for overall survival in oral cavity [pooled HR, 1.48; 95% confidence interval, (95% CI), 1.03-2.11], and for disease-free survival in oral cavity (pooled HR, 1.47; 95% CI, 1.12-1.93) and in oropharynx (pooled HR, 0.45; 95% CI, 0.27-0.73). Despite attempts to limit it, between-study heterogeneity was large in the majority of meta-analyses and the prognostic value of p53 was generally inconsistent and inconclusive across studies.

Conclusion: The meta-analysis results highlight that current evidence about the prognostic value of p53 in patients with squamous cell carcinoma of the head and neck is inconclusive. Large heterogeneity exists across studies in study-level and patient-level characteristics, making it difficult to ascertain a clear picture. Future studies are required in which p53 expression is investigated in a more standardized and biologically informative manner. In particular, prospectively planned individual patient data meta-analyses are needed to establish the prognostic importance of p53 for specific subgroups of patients undergoing specific treatments.

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Introduction

Loss of function of the p53 tumor suppressor gene plays an important role in the development of cancer (1). p53 regulates the activity of pathways, which lead variably to cell cycle arrest, DNA repair, senescence, or apoptosis following exposure of cells to endogenous or exogenous cellular stresses (2). With respect to cancer, those stresses that cause DNA damage are of particular interest. p53 has been extensively investigated across the whole spectrum of human malignancy, with mutations of the p53 gene (TP53) found in over 50% of cancers (3). p53 gene status in tumors such as carcinomas of the bladder affects upon disease outcome—mutations are linked with poorer patient survival, and this is likely due, in part, to reduced susceptibility of such cells to apoptosis (4).

Head and neck cancer accounts for 3% of incident cancer in the United Kingdom, with squamous cell carcinoma (SCC) being the commonest histologic type. The larynx, oral cavity, oropharynx, and hypopharynx comprise the four main anatomic sites within the head and neck, accounting for 90% of SCC cases (5). Smoking, alcohol, diet, and presence of the human papilloma virus (HPV) have all been identified as etiologic factors in the development of SCC of the head and neck (SCCHN; refs. 6, 7). Whatever the etiology, p53 overexpression (which is usually indicative of the presence of mutant p53) is also commonly seen in SCCHN, which is to be expected given an overall cited UK mutation prevalence of 50.7% (5). However, despite this knowledge and despite the fact that the presence of p53 expression and its effect on prognosis in SCCHN has been widely investigated in the published literature, no consensus exists about its true clinical role
in determining key parameters such as treatment responses or prognostic value. Existing studies are often small, heterogeneous (e.g., with respect to the primary tumor site and outcomes assessed), and often conflicting in their results, making it difficult to assess the clinical benefit of p53. Moreover, in many studies, the method used (immunohistochemistry) is often performed in a manner that is inherently ineffective at interrogating p53 function. This is disappointing because superior strategies based on the same simple technology have been widely known for several years (8).

The potential limitations of undertaking such a review were previously shown by Kyzas et al. in 2005 (9). They showed that the association between p53 status and mortality varied widely depending on whether indexed, published, or unpublished data were included in the meta-analysis. Such variation ranged from a risk ratio (RR) of 1.38 [95% confidence interval (CI), 1.13-1.67; \( P = 0.001 \)], to 1.23 (95% CI, 1.03-1.47; \( P = 0.02 \)), and to 1.16 (95% CI, 0.99-1.35; \( P = 0.06 \)) when indexed, published, or unpublished data were considered, respectively. These authors examined reports of p53 status in SCCHN tumors, and their meta-analysis of the data highlighted some important limitations and potential problems in such studies of p53 status compared with clinical outcomes. Although Kyzas et al. (9) reported results for tumors of the larynx and oropharynx separately, they did not include tumors originating from the other two main anatomic subsites from which SCCHN can arise, namely the oral cavity and hypopharynx. Due to the heterogeneous biological behavior and response to treatment of tumors originating from different subsites (10), it is essential to consider tumors originating from each site separately to avoid such confounding factors in this kind of meta-analysis. Furthermore, Kyzas et al. (9) gave priority to data from immunohistochemical (IHC) analyses when both IHC and genetic data (from PCR/single-strand conformational polymorphism) were available and used 2-year overall survival (OS) as their outcome indicator. Although IHC is the most commonly used method to examine such phenotype/outcome questions and has the merit of permitting the largest number of studies to be compared using a single method for p53 analysis, when performed for p53 alone—as it routinely is—it is not a reliable indicator of p53 status (as discussed below). Similarly, although OS at 2 years has the advantage of being readily obtainable information, disease-free survival (DFS) is arguably a more relevant outcome in this context.

In this article, we report our results from a systematic review and meta-analysis of existing studies specifically to investigate whether p53 is a prognostic factor in SCCHN in relation to OS and DFS for each of the major primary tumor sites. This allows us to produce evidence-based results for the prognostic value of p53 in SCCHN, to consider how heterogeneity across studies affects the observed prognostic effect of p53, and to identify further research needs.

Materials and Methods

Search Strategy

Medline (1966-2007), the Cochrane Central Register of Controlled Trials (1993-2007), and Embase (1974-2007) were systematically searched for relevant articles. The search strategy can be seen in Table 1. Conference abstracts, dissertations, and theses were also reviewed for any additional studies. Search terms used were as follows:

- p53
- Tumor suppressor protein p53
- Palatal neoplasm, cancer, or carcinoma
- Tongue neoplasm, cancer, or carcinoma
- Laryngeal neoplasm, cancer, or carcinoma
- Hypopharyngeal neoplasm, cancer, or carcinoma
- Oropharyngeal neoplasm, cancer, or carcinoma
- Tonsillar neoplasm, cancer, or carcinoma
- Oral cavity neoplasm, cancer, or carcinoma
- Mouth neoplasm, cancer, or carcinoma
- Head and neck squamous cell carcinoma
- Neck metastasis

Inclusion and Exclusion Criteria

The results from these searches were reviewed by title and abstract by two independent reviewers. Studies not written in English were excluded. If studies seemed appropriate, the full manuscript was scrutinized and the study deemed “relevant” if it met the following inclusion criteria:

1. Primary tumors were biopsy-proven nontreated SCCHN from the larynx, oropharynx, hypopharynx, or oral cavity (cutaneous and lip SCC, nasopharyngeal carcinoma, and cases of tumor recurrence were excluded).
2. Treatment modalities used included radiotherapy, chemotherapy, surgery, or a combination of these.
3. Studies assessed p53 status in pretreatment tumor samples and compared this with survival outcomes, and sufficient data were available to allow meaningful statistical analysis.
4. Survival outcomes were of a minimum 2-y follow-up after primary treatment.

These strict inclusion criteria were used to limit heterogeneity across studies and thus facilitate more clinically meaningful meta-analysis results. Between-study heterogeneity is a known problem for meta-analysis of prognostic marker studies (11), and we wanted to limit this problem from the outset by focusing on clinically similar studies with good follow-up that measured p53 at the...
same pretreatment baseline point and allowed primary tumor sites to be distinguished.

Outcomes

The primary outcome measures used were OS and DFS. As defined in our study, OS is the measure of the proportion of people within a treatment group who are expected to be alive after a specified time. It takes into account death due to any cause—both related and unrelated to the cancer in question. DFS, on the other hand, measures the proportion of people among those treated for a cancer that will remain free of disease at a specified time after treatment.

Data Extraction

Data were extracted from the full-published article for each study and these were collated in a Microsoft Excel worksheet. Site and stage of tumor, primary treatment, and number of cases analyzed for p53 were all recorded. The method of p53 analysis used was categorized as IHC or mutational analysis. IHC data recorded included the antibody used for analysis and the cutoff level used to
dichotomize p53-positive (high expression) versus p53
negative (low or undetectable expression) cases. Mutu-
tional analysis data included method of analysis—sin-
gle-strand conformational polymorphism with direct
sequencing—and the specific exons investigated. Authors
of trials were contacted and asked to supply fur-
ter data wherever necessary.

The recommended summary statistics for meta-analy-
thesis of time-to-event data (OS and DFS) are the log(hazard
ratio) and its variance, which account for both the time it
takes for an event to occur, as well as censoring. As
many trials do not report this information directly (12),
appropriate data such as log-rank test P values were ex-
tracted to allow estimation of the log(hazard ratio) and
its variance using previously reported methods (13). For
the studies (n = 6) that reported P values as less than a
given significance (e.g., P < 0.05) rather than reporting
the exact P value, the value reported as less than was
used to calculate the estimated hazard ratio (HR) for that
study group. If only Kaplan Meier graphs were pub-
lished, the “Mouseeyes” image digitization and linear
separation program (14) was used to extract data directly
from the survival curves (15). For the purposes of this
review, only HRs from unadjusted analyses were in-
cluded; we did not use results from adjusted analyses (that
is, those that assessed p53 in a model that adjusted for
other prognostic factors) because there was clearly a
large heterogeneity across studies in the sets of adjust-
ment factors that were used, and thus, meta-analysis
would have been more difficult to interpret. Where rele-
vant extractable data were not present (e.g., an unadjust-
ed HR), the primary author was contacted requesting
this for the purposes of our analysis.

Results from studies by the same authors were closely
scrutinized to avoid repetition of data in the meta-analy-
sis. If results for distinct patient groups were presented in
the same article, for example when results were separat-
ed by stage or treatment type, these groups were each in-
cluded in the meta-analysis as separate entries. Three
authors verified the accuracy of data extraction from in-
cluded articles independently. We used a software (ver-
sion 3.0; 28 Sept, 2004) developed by Matthew Sydes
and Jayne Tierney of the MRC Clinical Trials Unit, Lon-
don, to undertake calculations to estimate the trial level
log(hazard ratio) and variance based on the summary da-
ta extracted from publications.

Meta-analysis

Primary Analyses. Trial level log(hazard ratios) and
their variances were pooled using an inverse variance-
weighted random effects meta-analysis and results pre-
sentened graphically in the form of a Forest Plot with
pooled HR and 95% CI. The meta-analysis was undertak-
en using Review Manager (RevMan) version 5.0 (Copen-
hagen: The Nordic Cochrane Centre, The Cochrane
Collaboration, 2008). A random-effects meta-analysis
here estimates and accounts for any between-study het-
rogeneity, estimates the average (“pooled”) HR across
studies and its CI, and provides a prediction interval
for the HR in a future p53 study (16). It thus provides
an important summary of the current evidence base for
p53, and we were able to apply it to both outcomes
(OS and DFS) separately for the primary tumor sites of
larynx, oral cavity, and oropharynx, and for just the OS
outcome for the hypopharynx.

In addition, in which sufficient data were available for
each primary site individually, meta-analysis for OS and
DFS for early and advanced stage disease was calculated
each site. The effect of between-study heterogeneity
in our meta-analyses was assessed by the I2 statistic
(17). I2 describes the proportion of total variation in
meta-analysis estimates, which is due to between-study
heterogeneity rather than sampling error and is measured
from 0% to 100%, with increasing I2 values indicating a
larger effect of between-study heterogeneity in the meta-
analysis.

Subgroup Analyses. To investigate the effect of study
design and stage of disease on the p53 results, we also
performed subgroup analyses grouping studies that (a)
prospectively or retrospectively collected data, and (b) in-
cluded patients with advanced stage only compared with
early stage only. A test for subgroup differences was cal-
culated as described by Deeks et al. (18).

Publication Bias

For those meta-analyses including 10 or more studies,
we assessed the possibility of publication bias by visually
assessing a funnel plot for asymmetry and by performing
Egger’s test for such asymmetry with a 10% significance
d level due to the low power of this test (19).

Results

Search Results

Our search strategy identified 1,548 articles from Em-
base, 948 articles from Medline, 16 articles from the Co-
chrane library of randomized control trials, 3 articles
from conference proceedings, and 109 articles from the
bibliographies of retrieved articles. Following deduplica-
tion, two reviewers independently screened the titles
and abstracts identified. They subsequently agreed that
269 studies should be retrieved for detailed review. Forty-one relevant studies were eligible for inclusion (the
other 228 were excluded for several reasons; Fig. 1) and
the characteristics of these studies can be seen below. No
unpublished work was included.

- Awwad; mixed stage larynx, 40 cases; oral cavity, 14
cases; oropharynx, 14 cases treated with radiothera-
py; retrospective IHC study, cutoff is any staining;
OS and DFS (20)
- Buyukbayram; mixed stage larynx, 60 cases; retro-
spective IHC study, 20% cutoff; OS and DFS (21)
- Caminero; mixed stage oropharynx, 106 cases trea-
ted with surgery with or without radiotherapy; ret-
spective IHC study, 10% cutoff; OS (22)
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**Details of Systematic Review**

- Chien; mixed stage hypopharynx, 58 cases treated with surgery; retrospective IHC study, 50% cutoff; OS (23)
- Cho; stage I/II oral cavity, 33 cases treated with surgery with or without radiotherapy; prospective IHC study, 50% cutoff; DFS (24)
- Chomchai; mixed stage larynx, 45 cases mixed treatment; retrospective mutation analysis exons (5-8); OS & DFS (25)
- Costa; oral cavity, 51 cases treated with chemoradiation or surgery; prospective IHC study, cutoff not given; DFS (26)
- de Aguiar; mixed stage oral cavity, 81 cases treated with surgery ± radiotherapy; retrospective IHC study, 5% cutoff; OS (27)
- Fan; mixed stage larynx, 109 cases mixed treatment; retrospective IHC study, 10% cutoff; OS and DFS (28)
- Frank; mixed stage hypopharynx, 43 cases mixed treatment; retrospective IHC study, strong (10%) cutoff; OS (29)
- Friedman; stage III/IV larynx, 69 cases mixed treatment; retrospective IHC study, 10% cutoff; OS (30)
- Friesland; stage III/IV oropharynx, 37 cases mixed treatment; retrospective IHC study, 5% cutoff; OS (31)
- Grabenbauer; mixed stage oropharynx, 84 cases treated with surgery; prospective IHC study, 10% cutoff; OS and DFS (32)
- Hiranuma; mixed stage oral cavity, 45 cases mixed treatment; prospective mutation analysis exons (5-8); OS (33)
- Hirvikoski; mixed stage larynx, 99 cases mixed treatment; retrospective IHC study, 20% cutoff; OS (34)
- Homma; mixed stage larynx, 59 cases treated with chemoradiation; prospective IHC study, 10% cutoff; OS (35)

Figure 1. Details of systematic review.
• Jackel; mixed stage larynx, 88 cases treated with surgery with or without radiotherapy; retrospective IHC study, cutoff any staining; OS (36)
• Jayasurya; mixed stage oral cavity, 316 cases mixed treatment; prospective IHC study, cutoff median staining; OS and DFS (37)
• Jeannon; mixed stage larynx, 60 cases treated with radiotherapy; retrospective IHC study, 25% cutoff; OS (38)
• Kozomara; mixed stage oral cavity, 42 cases treated with surgery & radiotherapy; retrospective IHC study, cutoff any staining; OS (39)
• Miyahara; mixed stage hypopharynx, 28 cases; IHC study, 5% cutoff; OS (40)
• Nidu; stage III/IV hypopharynx, 53 cases treated with radiotherapy and surgery; retrospective IHC study, 10% cutoff; OS (41)
• Narayana; stage I larynx, 102 cases treated with radiotherapy; retrospective IHC study, 10% cutoff; DFS (42)
• Obata; mixed stage oropharynx, 33 cases mixed treatment; IHC study, 10% cutoff; OS (43)
• Pai; stage I/II larynx, 86 cases treated with radiotherapy; retrospective IHC study, cutoff any staining; DFS (44)
• Parikh; stage I/II larynx, 111 cases treated with radiotherapy; retrospective IHC study, 20% cutoff; OS and DFS (45)
• Pruneri; mixed stage larynx, 82 cases treated with surgery; mutation analysis exons 5-8; OS & DFS (46)
• Pulkkinen; mixed stage larynx, 62 cases mixed treatment; retrospective IHC study, cutoff intermediate staining; OS and DFS (47)
• Rowley; mixed stage larynx, 35 cases treated with radiotherapy; retrospective IHC study, cutoff any staining; OS (48)
• Russo; stage III/IV larynx, 81 cases treated with surgery; prospective mutation analysis exons 5 to 8; OS and DFS (49)
• Shigyo; mixed stage larynx, 32 cases treated with radiotherapy or surgery; prospective IHC study, 10% cutoff; DFS (50)
• Siegelmann; oral cavity, 45 cases; retrospective IHC study, 20% cutoff; OS and DFS (51)
• Sittel; stage I/II larynx, 47 cases treated with laser surgery; retrospective IHC study, 10% cutoff; DFS (52)
• Tan; mixed larynx, 90 cases treated with radiotherapy; retrospective IHC study, cutoff is strong staining; OS (53)
• Tatemoto; mixed stage oral cavity, 150 cases treated with chemoradiation with or without surgery; IHC study, 10% cutoff; OS (54)
• Tsujii; oral cavity, 50 cases; IHC study, cutoff any staining; OS (55)
• Vielba; mixed surgery larynx, 62 cases mixed treatment, retrospective IHC study, 5% cutoff; OS and DFS (56)
• Yanamoto; mixed stage oral cavity, 69 cases treated with surgery, IHC study, 10% cutoff; OS (57)

Primary Meta-analysis Results

Forrest plots of the primary meta-analyses for the individual primary tumor sites can be seen in Figs. 2 and 3. Table 2 reports the average (pooled) HR and its 95% CI for each of the meta-analyses. HR values of >1 indicate a poorer prognosis for overexpressing or mutant p53 tumors. There was some evidence from the meta-analyses that p53 status may provide prognostic information. The pooled HR was statistically significant for OS (pooled HR, 1.48; 95% CI, 1.03–2.11) and DFS (pooled HR, 1.47; 95% CI, 1.12–1.93) in relation to oral cavity, for DFS in relation to oropharynx (pooled HR, 0.45; 95% CI, 0.27–0.73).

However, despite our attempts to limit between-study heterogeneity through our strict inclusion criteria and by separating results by tumor site, there was still a large between-study heterogeneity in the effect of p53 status on the majority of meta-analyses, with $I^2$ generally toward 50% or above (Table 2). This means that the underlying HR for upregulated/mutant p53 versus wild-type or low/undetectable p53 actually varies across studies. Only in the oropharynx and oral cavity DFS meta-analysis was $I^2$ equal to zero, such that the between-study variance was estimated as zero. Therefore, only in these latter analyses can the pooled HR and its 95% CI be considered a reasonably reliable summary of the evidence.

In all the other meta-analyses that exhibit heterogeneity, it is clinically and statistically more informative to focus on the estimated 95% prediction interval for the underlying HR in a future study (16); this tells us the range of possible HR values in a new study that chooses its own particular set of study characteristics (e.g., cutoff level, treatment, etc.). All the calculated 95% prediction intervals in these meta-analyses contained the null value of 1 (Table 2), indicating that there is no significant evidence that p53 status maintains prognostic value regardless of the between-study heterogeneity. For example, the prediction interval for OS in the oral cavity site is from 0.57 to 3.86, indicating that given the heterogeneity, we are not sure whether upregulated/mutant or wild-type, low/undetectable p53 values indicate a worse prognosis, or indeed whether p53 has prognostic value at all. Further research is clearly needed to establish more precisely how study-level and patient-level characteristics are causing heterogeneity, and how they are
Figure 2. A, hypopharynx overall survival forest plot. B, larynx overall survival forest plot. C, oral cavity overall survival forest plot. D, oropharynx overall survival forest plot.
modifying the prognostic importance of p53 (see Discussion).

### Publication Bias

Due to the small number of studies in most meta-analyses, it was only sensible to examine the potential for publication bias in the larynx OS and DFS meta-analyses, which both contained >10 studies. There was no clear evidence of funnel plot asymmetry for either OS or DFS, either visually or from Egger’s test (OS: \( P = 0.95 \); DFS: \( P = 0.61 \)), and thus, there was no clear evidence of publication bias.

### Subgroup Analyses for Retrospective versus Prospective Studies

Only for the larynx and oral cavity tumor sites were sufficient articles available to compare OS and DFS between retrospective and prospective studies (Supplementary Fig. S1; Table 2). The subgroup analyses for OS excluded two studies from the larynx analysis and three

![Figure 3](https://www.aacrjournals.org/cebp/article-pdf/19/2/581/1247745/cebp_a10-0616_07.pdf)

**Figure 3.** A, larynx disease free survival forest plot. B, oral cavity disease free survival forest plot. C, oropharynx disease free survival forest plot.
from the oral cavity analysis due to lack of information to classify studies as retrospective or prospective. A statistically significant difference between the pooled results of prospective and retrospective studies was identified for the larynx OS analysis \((P = 0.0005)\) with pooled results from three prospective studies, suggesting increased length of survival for p53-negative status \((HR, 1.87; 95\% CI, 1.00, 3.51)\), and pooled results from 13 retrospective studies, suggesting an opposite effect in favor of p53-positive status \((HR, 0.74; 95\% CI, 0.53-1.02)\). However, heterogeneity between study results still remains within each subgroup and results should be interpreted cautiously. Statistically significant differences between the two subgroup pooled results were not identified for the other analyses (Table 2).

**Subgroup Analyses for Advanced versus Early Stage**

Only for the larynx tumor site were sufficient articles available to compare OS between studies with early and advanced stage patients (Supplementary Fig. S2; Table 2). However, this subgroup analysis only includes 3 of the original 18 studies with the OS data available and should be interpreted very cautiously. A statistically significant difference between the pooled results of advanced and early-stage tumors was identified for the

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NOTE: A HR of <1 indicates survival advantage for positive p53 status. A HR of >1 indicates survival advantage for negative p53 status.

Abbreviations: df, degrees of freedom; NA, not applicable—the between-study heterogeneity was estimated as zero so the pooled HR and its 95% CI alone are a suitable summary here.
Systematic Review of p53 in Head and Neck Cancer

The prognostic value of p53 in SCCHN has long been debated. When considering SCCHN, there is ample evidence of differing biological behavior and response to treatment depending on the primary site (10). Consequently, in this article, we have systematically reviewed whether p53 status is a prognostic factor for DFS and OS in patients presenting with carcinomas of the four main primary head and neck tumor sites, and ultimately included 41 studies in our review. In doing so, it is important to note that we excluded a large number of articles, which reported p53 status with survival outcomes, but without separation of the individual tumor sites. One such recent study was a prospective, multicenter review of p53 status and survival data for SCCHN treated surgically (61). This group extracted DNA from 420 cases (from all head and neck sites) to perform mutational analysis of exons 2 to 11 for the presence of disruptive or nondisruptive mutations. They found evidence of an association between p53 mutation and survival, showing that p53 mutations are significantly associated with a shorter OS in SCCHN compared with wild-type cases. Furthermore, they found that disruptive p53 mutations have a stronger association with shortened survival than nondisruptive mutations. This study contained significant results from a well-designed analysis of p53, but did not comply with the inclusion criteria of our review as it combined tumors from all sites as well as recurrent tumors.

Discussion

The prognostic value of p53 in SCCHN has long been debated. When considering SCCHN, there is ample evidence of differing biological behavior and response to treatment depending on the primary site (10). Consequently, in this article, we have systematically reviewed whether p53 status is a prognostic factor for DFS and OS in patients presenting with carcinomas of the four main primary head and neck tumor sites, and ultimately included 41 studies in our review. In doing so, it is important to note that we excluded a large number of articles, which reported p53 status with survival outcomes, but without separation of the individual tumor sites. One such recent study was a prospective, multicenter review of p53 status and survival data for SCCHN treated surgically (61). This group extracted DNA from 420 cases (from all head and neck sites) to perform mutational analysis of exons 2 to 11 for the presence of disruptive or nondisruptive mutations. They found evidence of an association between p53 mutation and survival, showing that p53 mutations are significantly associated with a shorter OS in SCCHN compared with wild-type cases. Furthermore, they found that disruptive p53 mutations have a stronger association with shortened survival than nondisruptive mutations. This study contained significant results from a well-designed analysis of p53, but did not comply with the inclusion criteria of our review as it combined tumors from all sites as well as recurrent tumors.

Inconclusive Evidence

Our meta-analysis results highlight that the current evidence base about the prognostic value of p53 is inconclusive, and there is large heterogeneity observed across studies. For all but two of our meta-analysis, due to the heterogeneity, we could not be sure whether positive or negative p53 values indicate a worse prognosis, or indeed that p53 has prognostic value at all; the prediction intervals for the HR in a future study are all wide (Table 2) and contain the null value of 1. Only for DFS in relation to oropharynx and oral cavity were there no estimated heterogeneity and a significant meta-analysis result. Why there was no heterogeneity in these particular settings is unclear, but caution is perhaps advised, as there were only three and five studies in the analyses. Interestingly, for oropharynx, the meta-analysis indicates that p53 over-expression/mutation confers a survival advantage; in contrast for oral cavity, the meta-analysis indicates that p53 overexpression/mutation confers a survival disadvantage. Such differences might derive from underlying biological differences or from sampling issues as outlined above due to the small numbers involved in these studies. However, the relatively high incidence of HPV infection in oropharyngeal SCC (40-70% of oropharynx tumors are HPV positive; refs. 62, 63) provides the most obvious potential biological source of altered p53 outcomes because HPV E6 inactivates p53 by promoting MDM2-independent degradation of p53 (64). However, studies of oropharyngeal cancers have found the presence of HPV to be associated with good outcomes (65, 66) and thus it is not trivial to explain why HPV infection might contribute to reduced survival in the p53-low/negative population relative to the survival for p53-positive patients as the apparent HR association suggests.

However, the relationship between HPV presence and p53 expression is complex as evidenced in a recent study of oropharynx cancers in which it was found that HPV DNA was present in 16 of 19 cases displaying wild-type p53 expression as confirmed by analysis of exons 4 to 9 (67) in these tumors. Therefore, the distribution of p53-positive samples in an HPV-positive cohort seems abnormal with respect to p53 mutational status and may well contribute to the results obtained from our meta-analysis. Clearly, even when prognostic significance is evident, the direction of effect is not consistent across tumor subtypes, and thus, we obtain an apparently confusing picture about how p53 status is associated with outcome. Nevertheless, such results provide further support for the need to perform separate analyses of individual tumor sites.

The Problem of Heterogeneity

It is clear that heterogeneity across studies here is severely “muddying the water.” We tried to limit such heterogeneity at the onset of our review by focusing on particular tumor sites, including just those studies with an adequate follow-up (2 or more years), ensuring p53 was measured in the pretreatment tumor tissue and, where possible, looking at advanced and early stage of disease separately. Yet despite this, much heterogeneity remained, and this is likely caused by other factors such as type of treatment, p53 cutoff levels, and method of measuring p53. It was very difficult to examine or explain this heterogeneity in our meta-analyses due to the generally small number of studies, the poor reporting of information, the lack of patient-level data, as well as the variability in clinical characteristics across patients within studies.

For example, although all included articles have reported results clearly for a single primary site, the majority have grouped together tumors of all stages, which will increase heterogeneity as the two most important prognostic factors for survival in SCCHN are primary tumor size (T stage) and nodal status (N stage; ref. 68). We
attempted to control for this by performing subgroup analysis for early (stage I and II) and advanced (stage III and IV) disease whenever adequate data were reported; however, this was only possible on two occasions (Table 2). In addition, fewer than half of the included articles reported a single mode of treatment for all patients; the remainder include patients treated heterogeneously using combinations of radiotherapy, chemotherapy, or surgery. Although in general it is reasonable to accept that on a stage by stage basis for the four primary sites, the differing treatment types carry similar survival outcomes, we have been unable to control for the inevitable effect of this treatment variability on survival outcome. Other patient demographics such as age, smoking, and alcohol use may also increase outcome variability.

In addition, the method of analysis of p53 status varies considerably between articles. Although IHC is the commonest method reported in the included articles, the primary antibody used to detect p53 expression was not consistent. Because it is not possible to distinguish wild-type from mutant p53 by IHC, the commonest antibodies used (DO7, DO1, and Pab 1801) are unable to differentiate tumor in which there is a primary p53 mutation from those in which wild-type p53 is upregulated in response to genotoxic stress. Thus, although p53 expression is commonly assumed to be a surrogate for the detection of mutant p53, this is evidently not always the case (8). Similarly, although the absence of p53 staining by IHC is considered synonymous with a wild-type phenotype, such a staining pattern could also be explained by gene deletion (8). A further issue is that the reported cutoff point chosen to dichotomize p53 positivity versus p53 negativity varied from the presence of one positively stained cell to 50% positivity. Both extremes are difficult to justify from a biological perspective. Occasional positive nuclei might result from the normal upregulation of wild-type p53 in response to localized stress. Heterogeneous upregulation of demonstrably mutant p53 has frequently been shown (69), and thus, using a cutoff of 50% will exclude many mutant samples that display lower levels of p53 positivity. The result of using such cutoffs would be, on one hand, to include wild-type samples in a presumed mutant population and, on the other, to place mutant samples with the wild-type group. Consequently, both of these scenarios would be expected to result in diminishing differences between p53 “wild-type” and “mutant” groups.

Other Problems Encountered During Our Meta-analysis Approach

Other problems beside heterogeneity limited our systematic review and meta-analysis; one problem in particular is the poor reporting of individual studies. This has also been highlighted elsewhere for prognostic markers in neuroblastoma (12). Despite our best efforts, we were unable to obtain all the published and unpublished pertinent data necessary for a complete review of p53 in SCCHN. For example, several articles discussed the prognostic value of p53 without presenting any useable data (despite our attempts to contact corresponding authors, these data were not made available to us). Of those studies that were included, the majority did not explicitly report their outcomes as a HR, necessitating the indirect estimation of HR using the techniques previously described (13, 15). On occasions, we used the 5-year survival rate to estimate the number of events, as the actual number was not directly reported; however, we are aware that if there was a lot of censoring, this may be an overestimate of the true number of events. This issue highlights the problems we found when extracting results from the literature for our meta-analysis. In addition, the use of inexact P values to calculate HR—as was necessary in the case of several articles—is a conservative approach and is likely to underestimate the true value of the HR. Results were only included from articles in which unadjusted results were available; published multivariate results were adjusted for different factors across individual studies, including other tumor markers, T stage, and N stage. These results were poorly presented—those factors not found to be significant in the model used were not reported.

As a consequence of this, synthesis of these studies into a pooled analysis would be problematic and hence was not undertaken. Selective reporting in studies of prognostic indicators, particularly the presentation of data, outcomes, and analyses, which give a statistically significant and favorable result, is also a concern. As explained previously, this problem has been highlighted by Kyzas et al. (70) who analyzed 1,915 articles on cancer prognostic markers and found that 90% to 96% of published articles gave a positive prognosis for the factor being investigated. Future primary studies should adhere to the REporting recommendations for tumour MARKer prognostic studies (REMARK) guidelines (71), which provide details of good practice for reporting prognostic tumor marker studies. We did not assess the quality of the primary studies in our review, although Hayden et al. (72) have recently provided some suggestions on how to do this.

Future Research

Given the heterogeneity and the difficulty in extracting suitable results from individual studies, it is thus not surprising that the prognostic value of p53 remains unresolved in SCCHN. Similar conclusions have been found in meta-analyses of p53 in bladder cancer, in which it was concluded that “after 10 years of research (including over 10,000 patients), evidence is not sufficient to conclude whether changes in p53 act as markers of outcome in patients with bladder cancer” (73), and “from this analysis, it becomes evident that further retrospective investigations will not contribute to the solution of the problem and thus are obsolete. There is an obvious need for standardization of the assay procedure and the assessment of the specimens as well as for the initiation of a prospective multicenter trial to provide definite answers” (74).
strongly echo these statements for the future analysis of p53 in SCCHN.

Indeed, a gold-standard approach would be to now initiate large, prospectively planned meta-analyses of individual patient data. Such an approach has many scientific advantages to a situation in which large (and small) centers conduct studies on their own patients, but with variations in inclusion criteria, measurement techniques, analysis strategies, etc. Riley et al. (75) discuss the benefits of individual patient data (IPD) and a prospectively planned IPD meta-analysis. These include the following:

- Consistency in the inclusion and exclusion criteria across studies
- Use of up-to-date follow-up information—potentially longer-term than that used in the study publication
- An ability to estimate for missing or poorly reported outcomes and summary statistics across studies—thus potentially reducing the problem of selective within-study reporting
- More direct estimation of the HR in which previously only indirect estimates were available
- Standardization of statistical analysis across studies
- The production of adjusted estimates in relation to a consistent set of adjustment factors, that is, an assessment of the value of p53 in relation to other existing prognostic markers
- The use of a consistent cutoff point across studies
- The assessment of the benefits of using combinations of markers
- The assessment of specific subgroups of patients across studies
- A proper investigation of the causes of study-level and patient-level heterogeneity in the prognostic effect of p53
- The identification of those studies that contain the same or overlapping sets of patients

McShane et al. (76) state that “More importantly, the necessity of large, definitive prospective studies or prospectively planned meta-analyses for tumor marker research must be recognized.” To achieve this, research groups need to collaborate and the focus should be on prospective, large, coherent, and high-quality research. “Cultural changes will be required” (76), but the necessity of this type of collaboration has been recognized before by clinical trialists and the epidemiologic community, and is achievable. With respect to SCCHN, research groups should particularly endeavor to examine the role of p53 in tumors from distinct primary sites, in strictly defined patient groups in which tumor stage and the treatment modality used are not only taken into consideration but are also controlled for. In addition, it is essential that consensus about the optimum methods of p53 detection as well as the scoring parameters chosen be agreed. For p53, there is simply no biological justification for using cutoffs for positivity of single nuclei, or 50% of cells, nor is the median a sensible criterion for dichotomization of this biomarker. Combinations of prognostic markers (e.g., p53 status and MDM2 and p21) should be investigated, as they can have increased prognostic power that exceeds that of the individual markers themselves (8, 77). It is relatively simple in this manner to improve upon the typically low specificity for p53 status determination by IHC and at the same time to increase the sensitivity of this process. For example, Nenutil et al. (8) proposed a simple strategy that combines highly sensitive detection of p53, with monitoring of p53 target gene expression (MDM2 and/or p21). Such an approach will greatly diminish the false positive and false negative mutation inference rate (inevitable with cutoffs of 1 positive nuclei and 50% positivity, respectively) and is amenable to standardization in scoring. For example, cells expressing p53 but not expressing MDM2 and/or p21, regardless of the relative levels, likely harbor mutant p53, whereas coexpression of p53 and MDM2 would be likely to indicate the retention of a wild-type p53 allele. Although such analyses are not as precise as direct analysis of DNA, the use of biologically justifiable combinations of markers can, for relatively little extra effort and using a low-tech approach (that is, IHC), provide a great deal of information about pathway function, and because it is a pathway function that determines cellular behavior, it is this that future studies must aim to interrogate.

Disclosure of Potential Conflicts of Interest

No potential conflicts of interest were disclosed.

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A Systematic Review of p53 as a Prognostic Factor of Survival in Squamous Cell Carcinoma of the Four Main Anatomical Subsites of the Head and Neck

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